LOCAL CROWDING OUT IN CHINA

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Introduction Data Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions Conclusions
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Outline

- Motivation and takeaways
- 2 Empirical approach
- Selated literature
- Oata
- Sesults
 - City-level
 - Industry-level
 - Firm-level
- Onclusions

Introduction Data Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions
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Motivation

- China responded to the global financial crisis with a massive fiscal stimulus:
 - In November 2008, the Chinese government announced a RMB4 trillion (USD590 billion) stimulus package
 - Implementation was immediate and mostly channeled via local governments
 - In 2009, city-level debt increased by RMB1.7 trillion
 - Central government debt increased by RMB700 billion

- Chinese local government debt almost quadrupled as a fraction of GDP between 2006 and 2013
 - $\bullet~$ From 5.8% to 22% of GDP

Introduction D	/ata	Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions
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Motivation

- The stimulus package focused on investment
 - In 2009, the growth rate of fixed capital formation almost doubled with respect to the pre-crisis period
 - In that year, the contribution of fixed investment to GDP growth was close to 90% (Wen and Wu, 2014)
- The investment boom was mostly funded by debt in state-owned enterprises (SOEs):
 - In manufacturing SOEs, leverage rose from 57.5% in 2008Q1 to 61.5% in 2010Q1
 - In private manufacturing firms, it dropped from 59% to 57% (Wen and Wu, 2014)

Introduction Data Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions
000000000000000000000000000000000000000				

Takeaways

- At first glance, the stimulus was a resounding success:
 - China escaped the great recession
 - It became one of the main drivers of world growth after the crisis (Wen and Wu, 2014)
- BUT the massive post-crisis increase in local public debt crowded out private investment (this paper)
- As private manufacturing firms have much higher productivity than SOEs (Song, Storesletten and Zilibotti, 2011), this reallocation of investment is likely to reduce China's long-run growth potential

Introduction Data Debt-to-GDP ratio			Conclusions
Our Empirical App	roach		

- City-level regressions
 - Investment of domestic *private* manufacturing firms is lower in cities with high public debt
 - No such correlation for SOEs and foreign firms
 - IV regressions: first evidence of causality from public debt to private investment
- Industry-level regressions
 - Public debt reduces investment by *private* manufacturing firms in industries with greater external financial needs
- Firm-level regressions
 - Investment sensitivity to cash flow is higher for *private* firms in cities with high public debt
 - Joint estimation of investment sensitivity and likelihood of being a constrained firm

Introduction Data Debt-to-GDP ratio City-level Regressions Industry-level Regressions Scooo Coord Coor

- Our identification strategy hinges on three features of Chinese capital markets:
 - Geographic segmentation: local public debt issuance (bonds and bank loans) does not attract capital from elsewhere ⇒ market imbalance remains local
 - Interest rate ceilings: local imbalance does not trigger increases in interest rates ⇒ no offsetting response by local saving ⇒ local rationing
 - Local governments (and SOEs) preferred by banks: local rationing only affects the investment of (unconnected) local private firms
- **Overall prediction**: local public debt issuance triggers rationing of local private & domestic firms **only**
- Note: not foreign firms, which can access non-Chinese capital markets

Evidence on Segmentation I

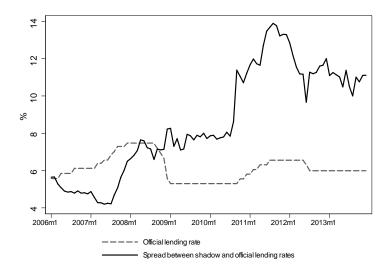
- Large banks are holding companies with separate legacy organizations for every province. A fragmented banking system in which local branches have substantial autonomy with respect to headquarters. Local politics and the pressure to lend to local governments and local state-owned enterprises play an important role (Dobson and Kayshap, 2006, p. 132)
- Local Communist Party officials have a bigger say in investment project approval than the head offices. Local authorities are crucial to bank managers' career advancement (Roach, 2006)
 - See also Boyreau-Debray and Wei (2004, 2005); Allen, Qian and Qian (2005); Brandt and Zhu (2007); Dollar and Wei (2007); Firth, Lin, Liu and Wong (2009).

Introduction Data Debt-to-GDP ratio City-level Regressions Industry-level Regressions occorrections occorrections

Evidence on Segmentation II

- We find that the interest rates of LGFV bonds at issue vary significantly and persistently between cities, controlling for credit rating and other bond characteristics. Municipal bond yields are positively correlated with local government debt
- Interest rate ceilings (on both deposits and loans) and the 2009 simulus package contrbuted to the growth of a shadow banking sector (from 14% of GDP in 2008 to 27% in 2010, Elliot, Kroeber, and Qiao, 2015, Chen, He, and Liu, 2016).
 - Jump in the spread between the shadow lending rate and the official lending rate
 - Entrusted loans between firms located in the same city carry a significantly lower interest rate than transactions between firms in different cities (Allen, Qian, Tu and Yu, 2016).

Official and Shadow Lending Rates



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Related Literature

- Public debt and growth:
 - Reinhart and Rogoff (2011), Cecchetti, Mohanty and Zampolli (2011)
- Effect of credit constraints on investment:
 - Fazzari, Hubbard and Pertersen (1998), Love (2003), Kaplan and Zingales (2000), Hu and Schiantarelli (1998), Almeida and Campello (2007)
- China:
 - Fiscal stimulus
 - Deng, Morck, Wu, and Yeung (2015), Ouyang and Peng (2015), Wen and Wu (2014), Cong and Ponticelli (2016), Bai, Hsieh, and Song (2016)
 - Capital misallocation
 - Song and Wu (2015), Hshei and Song (2016), Song, Storesletten, and Wu (2015)
 - Local government debt
 - Zhang and Barnett (2014), Wu (2015), Ang, Bai, and Zhou (2015), Ambrose, Deng and Wu (2015), Gao, Ru, and Tang (2016)

Introduction Data Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions
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Data

- 387,781 firms in 261 prefectural-level cities over 2007-13
- Sample covers 91% of Chinese population
 - City size ranges between 200,000 and 33 million
 - Median size: 3.8 million
 - 100 cities with a population greater than 5 million
 - 25 cities with a population greater than 8 million.
- Sources:
 - Firm level data
 - Annual Survey of Industrial Firms (ASIF) and Annual Tax Survey (ATS)

- City-level data (exlcuding local government debt)
 - China City Statistical Yearbook
- Local government debt
 - Part of our contribution: see next

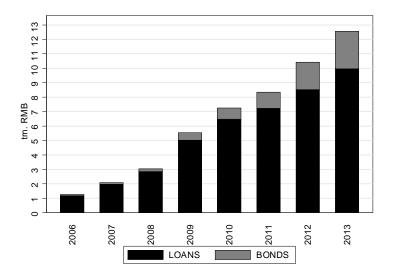


Local Government Debt

- We compute the *total* debt of Local Government Financing Vehicles (LGFV) by exploiting their reporting requirements
 - LGFVs that wish to issue a bond are required to disclose their current and previous balance sheets, at least for the previous 3 years
- We aggregate at city-year level and obtain a *lower bound* for total city-level public debt
 - But lower bound is much more accurate than data based on bond issuance only: see figures

• Wide cross-sectional and time-series variation in city-level public debt/GDP

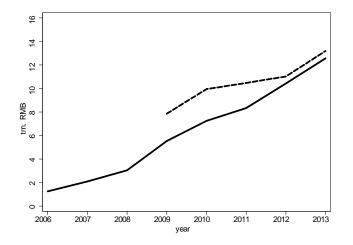
Composition of LGFV's Debt



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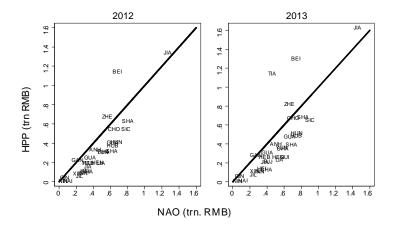
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Local Government Debt: Comparison with Official Data

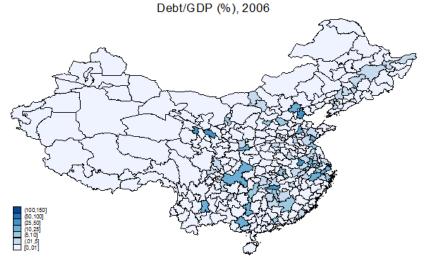


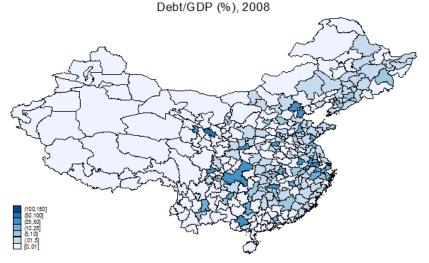
Solid line: our data; dashed line: official (CICC) data ヘロン 人間と 人間と 人間と Local Government Debt: Comparison with Province-Level NAO Data

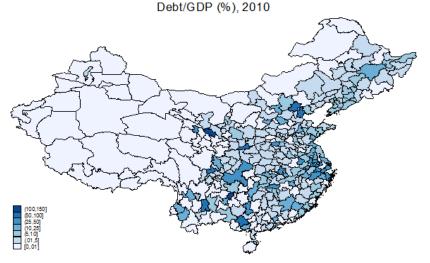
Introduction Data Debt-to-GDP ratio City-level Regressions Industry-level Regressions Firm-level Regressions Conclusions

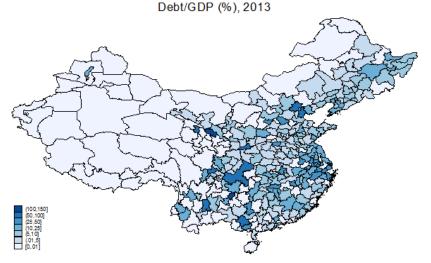


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Local Government Debt

Year	μ	σ	Min.	Max.	Max. Total China		Total China N. Cities	
		Bill.	RMB		Bill. RMB	(% GDP)	All	D>0
2006	4.3	18.1	0.0	173	1,255	5.8	293	92
2007	7.1	27.6	0.0	268	2,087	7.9	293	144
2008	10.4	38.4	0.0	383	3,036	9.7	293	189
2009	18.9	62.8	0.0	589	5,535	16.2	293	248
2010	24.7	80.5	0.0	789	7,249	18.1	293	281
2011	28.5	93.7	0.0	951	8,336	17.6	293	291
2012	35.6	113.0	0.0	1,145	10,425	20.1	293	292
2013	42.9	132.1	0.0	1,303	12,556	22.1	293	292

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Introduction Data Debt-to-GDP ratio City-level Regressions Industry-level Regressions Sociologo Conclusions Conclu

City-level Regressions

• We aggregate data at the city-year level and estimate the following regression:

$$I_{c,t} = \beta LGD_{c,t} + X_{c,t}\Gamma + \alpha_c + \tau_t + \varepsilon_{c,t},$$

- $I_{c,t}$ is the ratio of investment to the assets of manufacturing firms in city c and year t
- *LGD_{c,t}* is the ratio of local government debt to GDP in city *c* and year *t*,
- X_{c,t} are city-level controls
 - bank loans over GDP, local government balance over GDP, GDP growth, log of GDP per capita, log of population, and average land price

• α_c and τ_t are city and year fixed effects

City-level Regressions

	(1)	(2)	(3)	(4)	(5)
LGD	-0.083***	-0.089***	-0.017	0.017	
	(0.026)	(0.0289)	(0.029)	(0.052)	
LGD imes PRI					-0.090***
					(0.031)
LGD imes SOE					-0.029
					(0.028)
LGD imes FOR					0.0154
					(0.033)
N. Obs.	1,861	1,859	1,658	1,146	4580
N. Cities	261	261	261	245	261
Year FE	YES	YES	YES	YES	YES
City FE	YES	YES	YES	YES	YES
Sample	All	Private	State	Foreign	All
LGD × PRI –	-LGD imes SOE	(p-value=0.	01)		-0.060
LGD imes PRI -					-0.105
LGD imes SOE -	– $LGD imes FOI$	R (p-value=0	.13)		-0.045

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City-level Regressions: Additional Controls

(1)	(2)	(3)	(4)
0.093***	-0.104***	-0.029	0.032
(0.028)	(0.030)	(0.040)	(0.053)
-0.012	-0.002	-0.027	0.012
(0.014)	(0.014)	(0.024)	(0.033)
0.020	0.028	-0.139	-0.484*
(0.153)	(0.168)	(0.209)	(0.252)
).409***	0.332**	0.632***	-0.206
(0.127)	(0.135)	(0.164)	(0.190)
4.506	6.394*	-5.851	14.93**
(3.283)	(3.752)	(4.408)	(5.875)
7.506*	9.374**	-5.674	15.32**
(3.821)	(4.295)	(5.511)	(6.371)
0.598	0.505	-0.411	2.005*
(0.629)	(0.694)	(0.979)	(1.124)
1,805	1,803	1,658	1,109
261	261	261	242
All	Private	State	Foreign
	0.093*** (0.028) -0.012 (0.014) 0.020 (0.153) 0.409*** (0.127) 4.506 (3.283) 7.506* (3.821) 0.598 (0.629) 1,805 261	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	$\begin{array}{cccccccccccccccccccccccccccccccccccc$

All regressions include city and year FE

Introduction Data	a Debt-to-GDP ratio	City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions
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Endogeneity

- So far, these are just correlations
 - Local public debt may be endogenous, and the direction of the bias is not obvious
- Instrument local government debt with number of top national politicians born in a given city
 - "Connected cities" have more leeway to issue debt (Shih, Adolph and Liu, 2012, and Zhu, 2014)
 - They may also be deemed to be safer borrowers (more likely to be bailed out)
- But "connected cities" may also receive more transfers
 - We augment the regression with transfers, but also transfers are endogenous
 - We use a simulated instrument for transfers

$$STR_{c,t} = \frac{TR_{c,2005}}{TT_{2005}}TT_t$$

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City-level IV Regressions

Second Stage											
	(1)		2)	(3)	(4)				
LGD	-0.7	89**	-0.7	79**	-0.	446	-0.1	210			
	(0.3	368)	(0.3	383)	(0.3	310)	(0.2	277)			
TRI	0.4	54*	0.4	67*	0.0	883	-0.	131			
	(0.2	258)	(0.2	272)	(0.2	258)	(0.2	244)			
				First Stage							
	(1.1)	(1.2)	(2.1)	(2.2)	(3.1)	(3.2)	(4.1)	(4.2)			
	LGD	TRI	LGD	TRI	LGD	TRI	LGD	TRI			
TOP	0.13	2.48***	0.12	2.49***	0.03	2.75***	-0.23	3.11***			
	(0.41)	(0.81)	(0.4)	(0.82)	(0.44)	(0.89)	(0.43)	(1.02)			
STRI	0.39***	0.27	0.39***	0.28	0.40***	0.27	0.40***	0.23			
	(0.07)	(0.25)	(0.07)	(0.24)	(0.08)	(0.26)	(0.08)	(0.27)			
N. Obs.	1,8	361	1,8	359	1,!	575	1,1	127			
N. Cities		51	261		261		226				
CD F test	11	.44	11	.93	11	.92	12	.66			
Sample	A	11	Pri	vate	St	ate	Foreign				

All regressions include city and year FE

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Identification through Heteroskedasticity

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
LGD	-0.775**	-0.537**	-0.764**	-0.517*	-0.445	-0.445	-0.208	-0.0738
	(0.363)	(0.249)	(0.378)	(0.265)	(0.309)	(0.273)	(0.275)	(0.205)
TRI	0.453*	0.337*	0.466*	0.349	0.0888	0.183	-0.133	-0.0586
	(0.257)	(0.196)	(0.271)	(0.214)	(0.258)	(0.227)	(0.243)	(0.266)
EXT	2.488*	2.130*	2.581*	2.224*	0.406	0.0984	1.088	-0.304
	(1.353)	(1.249)	(1.428)	(1.326)	(2.200)	(2.261)	(2.786)	(2.295)
N. Obs	1,861	1,861	1,859	1,859	1,575	1,575	1,127	1,127
N. Cities	261	261	261	261	237	237	226	226
F test	11.6	10.4	11.7	10.5	11.99	11.05	13.04	11.68
Sargan test (p value)		0.51		0.54		0.81		0.69
Est.	IV	IV IH	IV	IV IH	IV	IV IH	IV	IV IH
City FE	YES	YES	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES	YES	YES
Sample	А	All	Priv	ate	St	ate	For	eign

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Firms with Low Exposure to Government Expenditure

	(1)	(2)	(3)	(4)	(5)
LGD	-0.089***	-0.103***	-0.938*	-0.932*	-0.653**
	(0.034)	(0.0378)	(0.502)	(0.498)	(0.304)
BL		-0.011			
		(0.018)			
GB		0.048			
		(0.205)			
GR		0.292*			
. ((0.154)			
In(GDP PC)		7.857*			
(000)		(4.645)			
In(POP)		7.571*			
		(4.381)			
LP		1.712*			
70		(0.929)	0 700**	0.000**	0 5 6 2 * *
TR			0.700**	0.699**	0.563**
EXT			(0.342)	(0.342) 0.879	(0.239) 0.284
EAT				(1.725)	(1.453)
N. Obs.	1,820	1,764	1,820	1,820	1,820
N. Cities	261	261	261	261	261
F test	201	201	11.4	11.6	10.9
J test (p value)			11.4	11.0	0.47
	LSDV	LSDV	IV	IV	-
Est.	LSDV	LSDV	IV	IV	IV IH

All regressions include city and year FE

Industry-level Regressions

• We use an approach akin to Rajan and Zingales (1998) to assess whether government debt affects more investment in industries that need more external funds:

$$I_{j,c,t} = \beta I_{j,c,t-1} + \delta \left(EF_j \times LGD_{c,t} \right) + \alpha_{j,t} + \theta_{c,t} + \varepsilon_{j,c,t},$$

- $I_{j,c,t}$: investment-asset ratio in industry j, city c and year t
- LGD_{c,t}: local government debt scaled by GDP in city c and year t
- *EF_j*: time-invariant measure of the external fund dependence of industry *j*
 - We build EF using data from the 4 Chinese cities with the most developed credit markets (Beijing, Shanghai, Hangzhou, and Wenzhou) and drop these 4 cities from the sample
- $\alpha_{j,t}$ and $\theta_{c,t}$ are industry-year and city-year fixed effects.

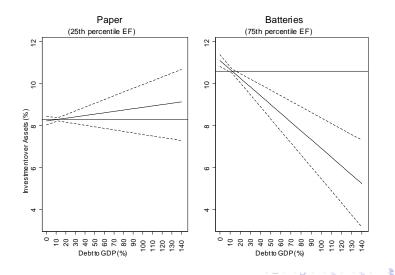
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Industry-level Regressions: Baseline

	(1)	(2)	(3)	(4)
I_{t-1}	-0.273***	-0.271***	-0.426***	-0.396**
	(0.006)	(0.006)	(0.034)	(0.16)
EF imes LGD	-0.015***	-0.019***	0.016	0.007
	(0.005)	(0.006)	(0.017)	(0.042)
N. Obs	57,054	53,262	6,249	2,550
N. Cities	15,768	14,906	3,252	1,121
City-Year FE	YES	YES	YES	YES
IndYear FE	YES	YES	YES	YES
Sample	All	Private	State	Foreign

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Introduction Data Debt-to-GDP ratio City-level Regressions Conclusions Coord Coord



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Industry-level Regressions: Additional Interactions

	(1)	(2)	(3)	(4)
I_{t-1}	-0.272***	-0.271***	-0.427***	-0.398***
	(0.006)	(0.006)	(0.03)	(0.164)
EF imes LGD	-0.018***	-0.023***	0.018	0.008
	(0.005)	(0.006)	(0.011)	(0.04)
EF imes BL	0.001	0.001	-0.003	-0.003
	(0.001)	(0.001)	(0.003)	(0.016)
$EF \times ln(GDP \ PC)$	0.227	0.186	0.679	-0.382
	(0.19)	(0.196)	(0.942)	(3.08)
EF imes GR	0.0286*	0.0338	0.0646	0.0191
	(0.016)	(0.019)	(0.09)	(0.312)
EF imes LP	-0.129	-0.131	-0.230	0.018
	(0.107)	(0.114)	(0.528)	(1.443)
N. Obs	56,209	52,503	6,065	2,520
N. Cities	15,693	14,839	3,194	1,115
City-Year FE	YES	YES	YES	YES
IndYear FE	YES	YES	YES	YES
Sample	All	Private	State	Foreign

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Firm-level Regressions

• We follow Fazzari et al. (1988) and Love (2003) and test whether local public debt affects the sensitivity of investment to internal cash flow:

$$I_{i,c,t} = \beta I_{i,c,t-1} + \delta REV_{i,c,t-1} + (\gamma_1 + \gamma_2 LGD_{c,t}) CF_{i,c,t-1} + \alpha_i + \theta_{ct} + \varepsilon_{i,c,t},$$

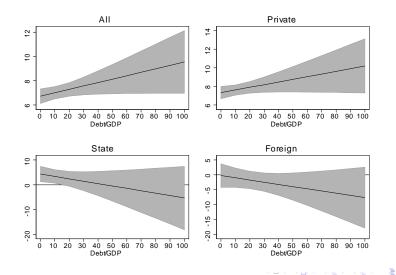
- *I*, *REV*, and *CF* are investment in fixed capital, revenue growth and cash flow of firm *i*, in city *c* and year *t* (all scaled by beginning-of-year total assets)
- *LGD* is local government debt scaled by GDP in city *c* and year *t*.
- α_i are firm fixed effects and θ_{ct} are city-year effects

Firm-level Regressions: Baseline

	(1)	(2)	(2)	(4)	(E)
	(1)	(2)	(3)		(5)
I_{t-1}	-0.273***	-0.280***	-0.371***	-0.282***	-0.273***
	(0.002)	(0.002)	(0.008)	(0.011)	(0.002)
REV_{t-1}	3.773***	3.799***	2.398***	2.942***	3.77***
	(0.031)	(0.034)	(0.167)	(0.220)	(0.031)
CF_{t-1}	6.725***	7.334***	4.328***	-0.253	6.70** [*]
	(0.231)	(0.256)	(1.190)	(1.534)	(0.231)
$CF_{t-1} \times LGD$	0.028**	0.029**	-0.097	-0.07	0.038***
	(0.011)	(0.013)	(0.055)	(0.05)	(0.012)
$CF_{t-1} \times LGD \times State$	(0.011)	(0.013)	(0.055)	(0.05)	-0.080**
$CP_{t-1} \times LGD \times State$					
					(0.036)
$CF_{t-1} \times LGD \times Foreign$					-0.091***
					(0.024)
N. Obs.	1,150,340	975,454	61,755	33,784	1,150,340
N. Firms	387,781	353,434	32,103	15,950	387,781
N. Cities	261	261	261	261	261
Firm FE	YES	YES	YES	YES	YES
City-Year FE	YES	YES	YES	YES	YES
Sample	All	Private	State	Foreign	All
$CF_{t-1} \times LGD + CF_{t-1} \times$			State	. s. cigii	-0.042
$cr_{t-1} \land cod + cr_{t-1} \land$ p-value	LOD A State				0.26
$CF_{t-1} \times LGD + CF_{t-1} \times$	LGD × Foreign				-0.053
p-value					0.11

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Introduction Data Debt-to-GDP ratio City-level Regressions Industry-level Regressions Conclusions Coordinate C



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Firm-level Regressions: Controlling for Bank Loans

	(1)	(2)	(3)	(4)
	(1)	(2)		()
I_{t-1}	-0.274***	-0.281***	-0.371***	-0.281***
	(0.002)	(0.002)	(0.008)	(0.011)
REV_{t-1}	3.770***	3.796***	2.393***	2.933***
	(0.031)	(0.033)	(0.168)	(0.220)
CF_{t-1}	8.343***	9.141***	6.020***	-2.973
	(0.374)	(0.411)	(1.893)	(2.665)
$CF_{t-1} \times LGD$	0.075***	0.083***	-0.045	-0.110*
	(0.014)	(0.016)	(0.068)	(0.058)
$CF_{t-1} imes BL$	-0.022***	-0.025***	-0.023	0.028
	(0.004)	(0.004)	(0.019)	(0.019)
N. Obs.	1,150,340	975,454	61,755	33,784
N. Firms	387,781	353,434	32,103	15,950
N. Cities	261	261	261	261
Firm FE	YES	YES	YES	YES
City-Year FE	YES	YES	YES	YES
Sample	All	Private	State	Foreign

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Firm-level Regressions: Additional Controls

	(4)	(0)	(0)	(1)	(=)
	(1)	(2)	(3)	(4)	(5)
I_{t-1}	-0.274***	-0.274***	-0.274***	-0.273***	-0.274***
	(0.002)	(0.002)	(0.002)	(0.002)	(0.002)
REV_{t-1}	3.771***	3.771***	3.796***	3.763***	3.787***
	(0.031)	(0.031)	(0.032)	(0.032)	(0.032)
CF_{t-1}	8.137***	9.150***	18.60***	2.039	19.15***
	(0.426)	(0.492)	(0.799)	(1.482)	(2.399)
$CF_{t-1} \times LGD$	0.075***	0.072***	0.052***	0.055***	0.051***
	(0.014)	(0.014)	(0.014)	(0.014)	(0.015)
$CF_{t-1} \times BL$	-0.021***	-0.024***	-0.026***	-0.025***	-0.021***
	(0.004)	(0.004)	(0.004)	(0.004)	(0.004)
$CF_{t-1} \times GB$	-0.038	· · ·	()	, ,	0.093*
	(0.042)				(0.052)
$CF_{t-1} \times In(GDP \ PC)$	()	0.539**			-0.794* [*] *
		(0.237)			(0.332)
$CF_{t-1} \times GR$		· · ·	-0.739***		-0.802***
. 1			(0.051)		(0.056)
$CF_{t-1} \times LP$			()	1.047***	-0.105
				(0.247)	(0.316)
N. Obs.	1,150,340	1,150,340	1,123,318	1,142,536	1,115,514
N. Firms	387,781	387,781	385,540	387,037	384,720
N. Cities	261	261	261	261	261
Firm FE	YES	YES	YES	YES	YES
City-Year FE	YES	YES	YES	YES	YES
Sample	All	All	All	All	All

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Firm-level Regressions: Exposure to Gov't Expenditure

	(1)	(2)	(3)	(4)	(5)	(6)
I_{t-1}	-0.277***	-0.278***	-0.283***	-0.375***	-0.304***	-0.278***
	(0.002)	(0.002)	(0.002)	(0.009)	(0.01)	(0.002)
REV_{t-1}	3.757***	3.756***	3.786***	2.368***	2.738***	3.756***
	(0.035)	(0.035)	(0.038)	(0.192)	(0.259)	(0.035)
CF_{t-1}	9.049***	8.455***	9.515***	7.913***	2.994	8.553***
	(0.442)	(0.421)	(0.487)	(2.360)	(3.410)	(0.477)
$CF_{t-1} \times LGD$	0.0895***	0.0785***	0.106***	0.029	-0.109	0.083***
	(0.0172)	(0.0156)	(0.020)	(0.079)	(0.086)	(0.020)
$CF_{t-1} \times BL$	-0.021***	-0.021***	-0.024***	-0.031	0.006	-0.021***
	(0.004)	(0.004)	(0.005)	(0.022)	(0.024)	(0.004)
$CF_{t-1} \times EXP$	-4.632***		-2.065*	-6.877***	-16.94	
	(1.009)		(1.236)	(2.128)	(11.24)	
$CF_{t-1} \times EXP \times LGD$	-0.064		-0.125**	-0.111	0.166	
	(0.046)		(0.052)	(0.105)	(0.481)	
$HEXP \times LGD$	-0.034**		-0.039**	-0.056	-0.071	
	(0.0136)		(0.0159)	(0.0384)	(0.0680)	
$CF_{t-1} \times HEXP$						-0.197
						(0.451)
$CF_{t-1} \times HEXP \times LGD$						-0.009
						(0.024)
$HEXP \times LGD$						0.003
						(0.004)
N. Obs.	935,255	935,255	796,947	50,192	24,087	935,255
N. Firms	323,914	323,914	295,448	26,065	11,790	323,914
N. Cities	261	261	261	261	261	261
Sample	All	All	Private	State	Foreign	All

All regressions control for firm and city-year FE

Firm-level System GMM Regressions

	(1)	(2)	(3)	(4)
	. ,	YS GMM	()	()
<i>I</i> _{t-1}	0.018	0.002	0.372	-0.404*
	(0.024)	(0.026)	(0.216)	(0.244)
REV_{t-1}	9.709***	9.756***	3.977	-0.607
	(0.365)	(0.407)	(3.882)	(3.494)
CF_{t-1}	9.69***	11.04***	36.15**	46.93*
	(2.41)	(2.69)	(17.48)	(22.80)
$CF_{t-1} \times LGD$	0.052***	0.037***	-0.044	0.056
	(0.011)	(0.012)	(0.046)	(0.123)
$CF_{t-1} \times BL$	-0.065***	-0.035	-0.066	-0.187*
	(0.020)	(0.023)	(0.106)	(0.170)
AR1 (p-value)	0.00	0.00	0.03	0.04
AR2 (p-value)	0.07	0.03	0.15	0.30
Sargan (p-value)	0.15	0.07	0.00	0.00
	Standard F	E on same sar	mple	
I_{t-1}	-0.242***	-0.251***	-0.339***	-0.206***
	(0.002)	(0.003)	(0.015)	(0.018)
REV_{t-1}	4.18** [*]	4.24** [*]	2.82***	1.07***
	(0.04)	(0.04)	(0.31)	(0.33)
CF_{t-1}	12.93***	12.87***	7.55**	15.32***
	(0.49)	(0.56)	(3.11)	(3.56)
$CF_{t-1} \times LGD$	0.018***	0.018***	0.005	0.021
	(0.002)	(0.002)	(0.013)	(0.013)
$CF_{t-1} \times BL$	-0.066***	-0.063***	-0.085***	-0.110***
	(0.005)	(0.006)	(0.030)	(0.027)
N. Obs.	797,314	623,837	53,657	18,848
N. Firms	261,451	190,525	19,136	6,028
Sample	All	Private	State	Foreign

All regressions control for firm and city-year FE

Firm-Level Regressions: Different Samples

	(1)	(2)
I_{t-1}	-0.282***	-0.278***
	(0.0018)	(0.0016)
REV_{t-1}	3.955***	3.793***
	(0.037)	(0.033)
CF_{t-1}	7.928***	8.352***
	(0.416)	(0.420)
$CF_{t-1} \times LGD$	0.057***	0.076***
	(0.019)	(0.017)
$CF_{t-1} imes BL$	-0.015***	-0.020***
	(0.004)	(0.004)
N. Obs.	781,670	1,003,337
N. Firms	264,914	340,510
N. Cities	235	212
Firm FE	YES	YES
City-Year FE	YES	YES
Sample	Excluding 4 provinces where HPP>Off.	1m <pop<10m< td=""></pop<10m<>

Local Government Debt and Investment: Firm-Level IV Regressions

(1)	(2)	(3)	(4)
-0.291***	-0.296***	-0.370***	-0.291***
(0.002)	(0.002)	(0.009)	(0.024)
3.659***	3.682***	2.358***	3.073***
(0.032)	(0.035)	(0.180)	(0.464)
23.65***	28.07***	20.08	2.736
(1.647)	(2.314)	(14.09)	(5.895)
2.638***	3.188***	2.176	1.829
(0.286)	(0.392)	(2.232)	(1.310)
-0.342***	-0.427***	-0.310	-0.154
(0.035)	(0.050)	(0.289)	(0.115)
-0.637***	-0.720***	-0.594	-0.824
(0.076)	(0.097)	(0.614)	(0.619)
928,772	775,250	43,617	19,130
261	261	256	2243
258,338	223,566	15,739	6,807
415.1	242.2	22.2	29.1
YES	YES	YES	YES
YES	YES	YES	YES
All	Private	State	Foreign
	-0.291*** (0.002) 3.659*** (0.032) 23.65*** (1.647) 2.638*** (0.286) -0.342*** (0.035) -0.637*** (0.076) 928.772 261 258.338 415.1 YES	-0.291*** -0.296*** (0.002) (0.002) 3.659*** 3.682*** (0.032) (0.035) 23.65*** 28.07*** (1.647) (2.314) 2.638*** 3.188*** (0.286) (0.392) -0.427*** -0.427*** (0.035) (0.050) -0.637*** -0.720*** (0.076) (0.097) 928,772 775,250 261 261 258,338 223,566 415.1 242.2 YES YES YES YES	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

Firm-Level Regressions: Switching Regression Model

- Two problems with an exogenous partition of firms across financing regimes (constrained or unconstrained):
 - Cannot jointly control for the various factors affecting how firms substitute internal with external funds
 - 2 Do not allow firms to switch across regimes over time
- As in Hu-Schiantarelli (1998) and Almeida-Campello (2007), we assume that at each date a firm operates in one of two regimes: a constrained regime where investment is very sensitive to internal funds, and an unconstrained regime where it is not
 - The regime probability is determined by a switching function that depends on firm characteristics that capture the severity of the agency problems faced by the firm at a specific date

Firm-Level Regressions: Switching Regression Model

• The regimes are not observable but endogenously determined by the system of equations:

$$W_{i,c,t}^* = M_{i,c,t}\psi + u_{i,c,t}$$

$$I_{1,i,c,t} = X_{i,c,t}\alpha_1 + \epsilon_{1,i,c,t}$$

$$I_{2,i,c,t} = X_{i,c,t}\alpha_2 + \epsilon_{2,i,c,t}$$

• A selection equation that estimates the likelihood that the firm is in regime 1 ($I_{i,c,t} = I_{1,i,c,t}$ if $W_{i,c,t}^* < 0$) or regime 2 ($I_{i,c,t} = I_{2,i,c,t}$ if $W_{i,c,t}^* \ge 0$) as a function of financial strength and other factors

Log of firm age, the log of total assets, distance to default -Altman Z-score-, industry-level
asset intangibility, firm ownership, local government debt

- Two investment equations
 - Identical to our baseline model but allowing for different coefficients in the two regimes.

Firm-Level Regressions: Switching Regression Model

	(1)	(2)	(3)
		A. Se	election Equation	1		
In(Age)	10.93	***	7.236	***	8.532	***
	(0.0	77)	(0.72	21)	(0.0	56)
In(Assets)	0.07	7**	0.725	***	1.706	***
	(0.0	34)	(0.03	30)	(0.0)	26)
Zscore	0.110)***	0.049	***	0.033	***
	(0.0	08)	(0.00	08)	(0.0	07)
Private	-9.34)***	-5.09	***	-4.339)***
	(0.1	42)	(0.0)	13)	(0.0)	12)
Tangible	7.898	***	4.62*	***		-
	(0.2	79)	(0.02	26)		
LGD	-0.0	12*				
	(0.0	08)				
N. Obs	1,060	,404	1,060	,404	1,060	,404
		B. Inv	estment Equatio	n		
	(1.1)	(1.2)	(2.1)	(2.2)	(3.1)	(3.2)
	Not Constr.	Constr.	Not Constr.	Constr.	Not Constr.	Constr.
CF_{t-1}	1.62***	0.40***	0.31***	0.81***	0.14***	0.71***
	(0.03)	(0.02)	(0.03)	(0.02)	(0.03)	(0.02)
$CF_{t-1} \times LGD$	-0.042***	0.014***	-0.063***	0.052***	-0.033***	0.011***
	(0.005)	(0.003)	(0.01)	(0.01)	(0.01)	(0.004)
LGD	-0.012***	-0.041***				
	(0.001)	(0.004)				
N. Obs.	306,175	754,229	274,822	785,222	231,925	828,479
City FE	YE		NO		NO	
Year FE	YE		NC		NO	
City-Year FE	N		YE		YE	
Ind-Year FE	N	C	NC)	YE	S

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Introduction Data Debt-to-GDP r	atio City-level Regressions	Industry-level Regressions	Firm-level Regressions	Conclusions

Conclusions

- In China local public debt issuance in 2007-13 crowded out the investment of private firms by tigtening their credit constraints
- We test this hypothesis owing to a novel dataset of city-level public debt, featuring strong cross-sectional and time-series variation
- Our evidence allays most endogeneity concerns
- The Chinese fiscal stimulus reduced investment by private manufacturing firms, and left SOEs unaffected
 - Reallocation of investment from private to SOEs is likely to reduce on China's long-run growth potential
 - By boosting the fraction of public debt in banks' portfolios, this policy has strengthened the bank-sovereign nexus in China, with possible systemic risk consequences: see Europe in 2010-12!