Exploring the Link between the Macroeconomic and Financial Cycles

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1. Introduction

The financial crisis prompted an increased focus on the role of financial factors driving real economic fluctuations. Recent papers have emphasised the significant economic effect of boom-bust cycles in credit and the overall importance of credit to the macroeconomy, including Reinhart and Rogoff (2009), Claessens, Kose and Terrones (2011), Borio (2012) and Aikman, Haldane and Nelson (2015). Several of these papers have suggested that there is a common cycle in financial variables (the financial cycle) that has a lower frequency and larger amplitude than the common cycle in real macroeconomic variables (the macroeconomic or business cycle), particularly since the 1980s. Borio (2012) defines the financial cycle as changes in market participants' perceptions of risk and attitude towards risk as well as changes in financing constraints, all of which tend to reinforce one another and lead to booms and busts in credit and asset prices, with the potential for these busts to have serious implications for financial stability, and therefore the broader economy, over a significant length of time. As a result of this work, the concept of a financial cycle that is different from the macroeconomic cycle has become more widespread within the academic and regulatory communities. There has been increased attention on policies that aim to manage the financial cycle; in particular, the focus has been on tempering the build-up in financial stability risks that often occur in phases when rapid asset price inflation is being fuelled by high credit growth, and implementing policies that encourage the financial system to support economic activity following a financial crisis.[1]

Three questions naturally arise from the literature's findings: why might the financial cycle be different from the macroeconomic cycle; does the financial cycle operate at a lower frequency than the business cycle; and how might these cycles be related? These questions are important for macroeconomic policymakers since the synchronisation and relationship between the cycles may influence the decision of whether to manage the business cycle separate from developments in the financial system or whether these developments (and any policies targeting these developments) should be considered when setting macroeconomic policy.

Since the fundamental role of the financial system is to facilitate the reallocation of resources from savers to investors, economic conditions should affect the resources available in the financial system. Economic conditions should also affect the demand for the financial system's services and borrowers' ability to repay their debts. As a result, it might be expected that the cycle in financial variables would be largely the same as the cycle in macroeconomic variables, abstracting from any phase shifts. However, it is likely to be more complicated than this, with other potential factors playing a role. The literature provides a number of reasons for why the financial cycle might differ from the macroeconomic cycle, including: the presence of financial frictions (Bernanke and Gertler 1989; Kiyokaki and Moore 1997; Adrian and Shin 2010); the tendency of financial market participants to inappropriately respond to risk (Borio, Furfine and Lowe 2001); and changes to policy regimes that can allow credit and asset price booms to take place at the same time as inflation (and the policy rate) remains low (Drehmann, Borio and Tsatsaronis 2012).

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1   For example, the Bank for International Settlements' countercyclical capital buffer and the Bank of England's new stress testing framework.

Developments in the financial system can also have an important effect on economic conditions. In particular, the financial system's resource reallocation role enables investment to take place that would have otherwise had to be self-financed, thereby boosting income and asset prices (Lowe 1992). The influence on asset prices can then affect household consumption and business investment through household net worth, business net worth and the replacement value of capital stock (Cochrane 2008). As a result, credit and asset price booms may be able to influence economic conditions and amplify economic booms and downturns. If the financial cycle does differ from the macroeconomic cycle and financial conditions can influence economic conditions, there is the potential for conflict between these cycles to have a real effect on the economy and at times, when these cycles are in different phases, may pose a conundrum for policymakers.

While there has been much progress made in this area, work is by no means complete. Given the importance of this work for policymakers, we believe that further work is needed to help improve the understanding of cycles in financial and macroeconomic variables and the relationship between these cycles. This paper adds to the existing literature by testing the differences in the cyclical properties of financial and macroeconomic variables, examining the extent to which the financial cycle is linked with the macroeconomic cycle, and considering the implications of our results for monetary policy and financial stability policies. The remainder of the paper is structured as follows: the next section summarises the relevant literature; the third section outlines the theory, data, methodology and results; the fourth section provides discussion on the policy implications of the results; and the final section concludes.

2. Summary of the Literature

Much of the recent literature on financial cycles builds on the arguments put forward in papers written as early as the 1930s (Fisher 1933; Gurley and Shaw 1955; Minsky 1977; Kindleberger 1978). In contrast with the mainstream economic literature in the second half of the twentieth century, which argued that there was no role for finance in determining economic conditions (e.g. Modigliani and Miller 1958), these papers suggest that the financial system can propagate shocks as well as be the source of shocks that lead to economic instability. According to Minsky, financial crises are not random or the result of poor policy decisions. Instead, they are due to the normal functioning of the economy endogenously leading to fragile financial structures; in particular, the pro-cyclical supply of credit creates a fragile financial system and increases the probability of a crisis occurring. These papers tend to be based on conjecture and more theoretical in nature, which reflects the lack of data that was available at the time.

Over the following decades, this work continued to be overshadowed by the literature that separated finance from economics. In the 1990s and early 2000s, new research emerged focusing on the pro-cyclicality of the financial system and the concept of a 'financial cycle' was introduced. Bernanke and Gertler (1989), Kiyotaki and Moore (1997) and Bernanke *et al* (1999) develop theoretical frameworks that explicitly allowed for financial frictions to amplify and propagate shocks to the economy, for example due to the pro-cyclicality of borrower net worth; this is commonly referred to as the 'financial accelerator' mechanism. Relatedly, Fostel and Geanakoplos (2008), Adrian and Shin (2010) and Geanakoplos (2010) highlight the importance of collateral rates (or 'leverage') for the economy; in particular, the pro-cyclical variation in leverage can significantly affect asset prices, which can subsequently influence the behaviour of economic agents.

In addition to the financial accelerator mechanism, Borio *et al* (2001) argues that an additional source of financial pro-cyclicality is market participants' inappropriate responses to changes in risk, which can result in significant financial and economic instability. Reflecting this, the paper suggests that policymakers should consider promoting a better understanding of risk and ensure supervisory rules promote better measurement of risk by using supervisory instruments in a countercyclical manner. The authors propose the use of monetary policy to contain financial imbalances and are some of the first to discuss the concept of the 'financial cycle', which they define as the expansion and subsequent bust in credit and asset prices, accompanied by changes in lending standards. A related paper, Borio and Lowe (2002), finds credit growth and asset price inflation to be good predictors of financial crises, and notes that financial imbalances can build up in relatively weak economic conditions. As a

result, the authors argue that financial stability and economic conditions (and therefore the business cycle and financial cycle) can be in conflict in the short term, and therefore both monetary and prudential policymakers should be allowed to respond to financial imbalances in a weak economic environment.

Since the financial crisis there has been a resurgence of work in this area, with many papers building on the earlier work and looking to determine the specific characteristics of the financial cycle, how these differ from the characteristics of the business cycle, and the relationship between these cycles. Many of these papers motivate their work based on improving the understanding of these cycles and their interactions to help macroeconomic policymakers achieve their mandates.

Reinhart and Rogoff (2009) examines a range of different financial crises and conclude that a common theme among these crises is excessive debt accumulation, which suggests that strong credit growth can have significant implications for the financial system and economy. Schularick and Taylor (2012) and Jordà, Schularick and Taylor (2015) use cross-country panel data and find that, since the Second World War, booms in credit and housing prices have been good predictors of crises that lead to large effects on economic output, and should therefore be carefully monitored by policymakers. The authors believe that this mainly reflects failures in the operation and regulation of the financial system rather than credit booms *per se*, and assert that their results are consistent with the notion that the financial system itself can generate economic instability.

A variety of econometric techniques have been used to empirically test the characteristics of the business and financial cycles, with most papers focusing on cycles in advanced economies. The four main techniques that have been used include: turning point analysis; frequency-based filters; model-based filters; and spectral analysis.

- The papers that use the Bry-Boschan turning point algorithm (updated to quarterly data by Harding and Pagan 2002) find that recessions following credit crises and asset price busts are generally longer lasting and more costly, financial cycles tend to be longer than business cycles, and these cycles can be highly synchronised (Claessens *et al* 2011; Drehmann *et al* 2012). There are some drawbacks to using the turning point technique; in particular, there is little theory available or offered to help explain the results, and the results can be sensitive to censoring assumptions.

- The use of frequency-based filters to obtain cycles in financial variables and cycles in output provides evidence that financial cycles are longer in length and larger in amplitude than business cycles and are reasonably synchronised (Drehmann *et al* 2012; Aikman *et al* 2015; Schüler, Hiebert and Peltonen 2016). However, in most cases, these papers make *ex ante* assumptions about the length of a business cycle, resulting in longer financial cycles and shorter business cycles by construction.[2] Moreover, this work tends to disregard the research that finds output to have significant medium-term cycles (Comin and Gertler 2006). In addition, frequency-based filters can introduce spurious cycles into the results (Schenk-Hoppé 2001; Phillips and Jin 2015; Hamilton 2017).

- A few papers have opted to use model-based filters rather than frequency-based filters, which involves applying the Kalman filter to unobserved components models in order to extract cycles (Rünstler and Vlekke 2015; Galati *et al* 2015). These papers have found that financial variables exhibit medium-term cycles that are longer than the traditional length of the business cycle (typically considered to be between 2½ and 8 years), though they also find some countries have a business cycle longer than 8 years. Also, medium-term cycles in financial series are found to be highly correlated with the medium-term component in output. For the purpose of identifying cycles in data, a significant drawback of using this approach is its model dependence; the results depend on the assumptions made about the underlying data generating process.

- Several recent papers have looked at financial and economic variables in the frequency domain to understand the difference between the financial and business cycles (Aikman *et al* 2015; Strohsal, Proaño and Wolters 2015; Schüler *et al* 2016). These papers find evidence of a financial cycle that is longer than the business cycle and evidence for both short- and medium-term cycles in real loan growth. However, by only using the

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2   Schüler *et al* (2016) allow the frequency range to vary by country based on multivariate spectral density results, and so did not assume the length of the financial and business cycles.

frequency domain, spectral analysis does not identify a time series measure of the cycle (without the use of a frequency-based filter) and generally requires stationarity, implying that it can only identify growth cycles.[3]

These techniques could be influenced by structural breaks in the data, though this should be less of a concern for turning point analysis. For example, the results found using some of these techniques may be affected by a lengthening in the business cycle; this is a potential explanation for the results found in Aikman *et al* (2015) and would suggest that it is not appropriate to compare the medium-term cycles in financial variables with short-term cycles in economic variables.

Despite some drawbacks, all four techniques seem to produce similar results, with most advanced economies found to have a longer common cycle in financial variables than the cycle in output, and some relationship between these cycles. However, the existing literature lacks any robust testing on whether the financial and business cycles are different. In addition, apart from conjecture and observations from the data, it remains unclear how these two cycles relate to one another and what drives this relationship.

As a result, our paper contributes to this literature by: estimating the length of the financial and business cycles and providing confidence intervals around these estimates; analysing the relationship between these cycles; and providing some discussion around the policy implications of our results. To do this, we make use of the techniques commonly used in the literature in such a way that aims to minimise the effect of their drawbacks on our results.

3. Financial and Macroeconomic Cycles

3.1 Theoretical models

Before we present our empirical methodology, it is useful to outline some of the theoretical economic models that incorporate a financial sector, and what these theoretical models might imply for our results and optimal policy decisions. Under the Modigliani and Miller (1958) irrelevance theorem, real economic outcomes are not affected by financial conditions, thought this assumes that no financial frictions are present. Given that financial frictions do exist, several influential papers have outlined theoretical models that explore what a financial system with frictions looks like and how a financial system with financial frictions can affect macroeconomic conditions.

One of the earlier theoretical papers is Bernanke and Gertler (1989). The authors develop a simple dynamic general-equilibrium model to show how borrowers' balance sheets can affect macroeconomic conditions. In this model it is assumed that there is asymmetry of information between borrowers and lenders and therefore some deadweight loss will occur; that is, the agreed financial contract is suboptimal to the financial contract with no asymmetry of information. Bernanke and Gertler argue that this deadweight loss (or agency cost) varies inversely with borrower net worth. If net worth varies with the business cycle, agency costs are countercyclical and amplify swings in borrowing, and therefore also investment, consumption and output. If there is a shock to net worth that is independent of economic conditions, the financial system can be a source of real economic fluctuations. Overall, this model suggests that the financial system can *amplify* macroeconomic shocks (referred to as the 'financial accelerator' mechanism) as well as be a *source* of macroeconomic shocks.

Kiyotaki and Moore (1997) is another important paper in the literature. It uses a dynamic equilibrium model to show how credit constraints influence macroeconomic conditions. In contrast to the model in Bernanke and Gertler (1989), where changes in net worth are due to changes in cash flow, Kiyotaki and Moore use changes in asset prices to be the source of changes in borrower net worth. In their model, durable assets act as both a factor of production and as collateral. As a result, there is a dynamic relationship between credit constraints and asset prices; a temporary shock that affects asset prices reduces the net worth of constrained agents and leads to tighter borrowing constraints, which lowers production and spending and further reduces asset prices.

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3   Some papers, such as Priestley (1965) and Bayram and Baraniuk (2000), develop methods for estimating the spectral densities of non-stationary processes, referred to as time-varying power spectra.

Building on the these two theoretical models, Bernanke *et al* (1999) develops a framework for credit-market frictions that incorporates price stickiness, a role for monetary policy, decision lags in investment and heterogeneity among firms. These additions aim to incorporate the results of the empirical literature and improve the model's empirical relevance. Similar to the previous literature, the authors found that the financial accelerator has an important effect on macroeconomic conditions. In this model, the role of the financial accelerator mechanism in amplifying the business cycle would be muted to the extent that monetary policy could stabilise output, though changes in policy have to be quite smooth in order for output to be stabilised.

Further building on the existing work and incorporating lessons learnt from the financial crisis, Gertler and Karadi (2011) and Gertler and Kiyotaki (2011) develop frameworks that allow for a significant disruption in financial intermediation. These papers not only consider how such a disruption would be expected to affect macroeconomic conditions, but also how policy responses, such as credit market interventions, might mitigate these effects. Gertler and Karadi introduce shocks to the asset quality of financial intermediaries, while Gertler and Kiyotaki introduce disruptions to financial intermediaries through constraints to obtain both depositor and wholesale funds, both of which work to reduce credit, increase the cost of credit, and therefore reduce economic activity. Both papers find that the net benefits of intervention in the credit market can be substantial, though this does depend on the severity of a crisis and the efficiency costs of government intermediation.

The recent crisis also increased the focus on the effect that leverage can have on the macroeconomy. In particular, Geanokoplos (2003; 2010) highlight the procylicality of collateral rates due to uncertainty around future asset price growth and the subsequent effect this has asset prices. The model in these papers assumes heterogeneity among buyers; for some reason one group of buyers values the asset more than any other buyers and are therefore willing to pay more for it, which often involves taking on more leverage. If there is a shock to these highly leveraged buyers, either from collateral constraints or a balance sheet shock, they will buy less and the asset price will fall. This shock is often brought on by bad news that increases uncertainty of future asset returns, causing lenders to increase margins. As a result, a feedback loop will emerge that leads to sharper declines in asset prices. Geanakoplos argues that, in absence of intervention, leverage becomes too high in a boom and too low in a downturn, and therefore the central bank should consider restricting leverage in boom periods to avoid future financial crises. Fostel and Geanakoplos (2008) uses a similar model to show that leverage cycles in an economy that has experienced some bad news (but not enough to warrant panic) can lead to contagion, a flight to well-collateralised assets and bond issuance rationing.

While much of the literature focuses on secured credit, Azariadis, Kaas and Wen (2015) observe that unsecured credit shows greater cyclicality than secured credit and often leads GDP, challenging the notion that it is collateralised debt that plays a key role in influencing macroeconomic conditions. To explain this result, the authors develop a dynamic general equilibrium model in which the unsecured component of the loan depends on a firm's expectations of the future availability of unsecured credit, since these expectations represent the firm's chance of remaining solvent in the future. In addition to these constraints depending on future expectations of credit conditions, credit constraints and productivity are endogenous in the model; when constraints are binding, productivity falls causing constraints to rise. As a result, the model gives rise to self-fulfilling credit cycles, where shocks to expected future unsecured credit conditions can have persistent effects on credit, productivity and output.

The theoretical literature suggests that the financial system can be influenced by factors that can be independent of economic conditions and emphasises the link between the macroeconomy and the financial system. This suggests that the financial cycle could be different from the business cycle, though these cycles should still be connected even if they are different in length. What does this mean for policymakers? The current work suggests that policymaking that only targets macroeconomic conditions and does not take developments in financial system into account may not effectively prevent the occurrence of financial crises and the subsequent negative effects on the macroeconomy. We will come back to the discussion on policy implications in Section 4.

3.2 Data and descriptive statistics

We use several financial and macroeconomic variables to obtain indicators of the financial and macroeconomic cycles: the financial variables are credit, housing prices, equity prices and 10-year government bond rates; and the macroeconomic variables are gross domestic product (GDP), employment and the unemployment rate.[4] All nominal variables are deflated using the consumer price index. These variables are generally consistent with the variables used in the literature, though some papers have chosen only credit and housing prices in their measures of the financial cycle. Given that there is some evidence that other financial variables could affect systemic risk (Jordà *et al* 2015), we choose to include all four measures in our analysis and perform sensitivity analysis to determine whether our results are robust to the exclusion of equity prices and government bond yields. It should be noted that we do not aim to determine which variables best represent the financial cycle; instead we choose widely available metrics that capture activity in the financial system.[5]

For parts of our analysis, we only use credit growth and GDP growth to represent the financial and macroeconomic cycles, respectively. This approach is similar to several other papers, including Strohsal *et al* (2015) and Aikman *et al* (2015). We have chosen to do this for simplicity; we acknowledge that it is not ideal given that GDP does not always capture the true state of an economy and the interaction between house prices and credit is important (Drehmann *et al* 2012; Jordà *et al* 2015).

For most of our analysis, the variables are expressed as log annual differences, except for unemployment rate and 10-year government bond yields which are expressed as annual differences. These data transformations are made to induce stationarity in the series, since most of our analysis requires the data to be stationary.[6] Five countries are included in our analysis: the United States, the United Kingdom, Australia, France and Germany. These countries were chosen to facilitate a comparison with the findings in the literature as well as provide a comparison across advanced economies.

A summary of the data is presented in Table A1, and the data are shown in Figures A1–A5 in Appendix A. The summary statistics are broadly consistent across countries, which is not surprising given the similarities between the countries. However, Germany appears to be somewhat of an outlier, with negative average house price inflation and weaker credit growth over the past couple of decades. On visual inspection, credit growth and housing price inflation generally appear to have had larger and longer cycles compared with equity price inflation and the change in government bond yields, which is a common finding in the literature (Drehmann *et al* 2012, Strohsal *et al* 2015, Schüler *et al* 2016). There also appear to be cycles in the macroeconomic variables.

3.3 Methodology

As discussed in Section 2, several papers approach the question of whether there is a longer financial cycle by imposing *ex ante* assumptions on the cycles in the data. We use two methods that do not impose such assumptions. The first method is multivariate spectral analysis, which decomposes the cross-covariance of stationary time series into different cycle frequencies. The second method is the turning point algorithm developed in Bry and Boschan (1971), which has been extended to quarterly data in Harding and Pagan (2002) and multiple variables in Harding and Pagan (2006). It is worth noting that, like many existing papers, we estimate these cycles separately and as a result we are treating the cycles as if they are separate from each other. If they are related, a better approach might be to estimate the cycles simultaneously; however, we leave this for further work.

For the first method, the frequency with the greatest contribution to the cross-covariance is considered to be the dominant cycle length at which the time series co-move together. The dominant cycle in the financial time series is

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4   See Appendix A for data sources and sample periods.
5   While the credit gap and credit-to-GDP gap have been commonly used to measure the financial cycle, there is little theory behind these measures. In contrast to the output gap, it is not clear what the level of potential credit is or what distortion makes credit deviate from its potential.
6   We use the variables in level terms for turning point analysis.

used as an approximation of the financial cycle, and the dominant cycle in real economic time series is used as an approximation of the macroeconomic cycle. To do this, we follow Schüler *et al* (2016) and calculate a measure of the multivariate spectral density called the Power Cohesion:

$$PCoh_X(\omega) = \frac{1}{(M-1)M} \sum_{a \neq b} |f_{x_a x_b}(\omega)|$$

$$f_{x_a x_b}(\omega) = \frac{s_{x_a x_b}(\omega)}{\sigma_{x_a} \sigma_{x_b}} = \frac{1}{2\pi} \sum_{k=-\infty}^{\infty} \frac{Cov[x_{a,t}, x_{b,t+k}]}{\sigma_{x_a} \sigma_{x_b}} e^{-ik\omega}$$

where t = 1,....,T reflects the time dimension, M reflects the number of stationary variables in the matrix X, which is a T x M matrix, and $\omega \epsilon [-\pi, \pi]$ reflects the cycle frequency. The function $f_{x_a x_b}(\omega)$ represents the cross-spectral densities for each pair of variables in X, with σ_{x_a} and σ_{x_b} representing the standard deviations variables *a* and *b*, $s_{x_a x_b}(\omega)$ representing the cross-spectrum for variables *a* and *b*, and $Cov[x_{a,t}, x_{b,t+k}]$ representing the cross-covariance between these variables; *i* is the imaginary unit. In our analysis, we have two X matrices, one for financial variables and another for economic variables. The dominant cycle is determined by finding the frequency, ω, at which $PCoh_X(\omega)$ is at its maximum. To determine whether these dominant cycles are different, we estimate a 95 per cent confidence interval for the dominant cycle frequencies using a parametric bootstrap method.[7]

There are two main limitations with using multivariate spectral analysis. Due to the volatility in the multivariate spectral density, we use a Parzen window to smooth the cross-spectral densities; while this could potentially compromise our results, we note that it is commonly used in the literature (e.g. in Aikman *et al* 2014 and Scüuler *et al* 2016) and changing the bandwidth (or removing the smoothing window altogether) does not change our main conclusions. Another limitation with this method is that it requires stationarity and therefore we are only able to identify cycles in growth rates rather than classical cycles. It could also be a concern that we include both stock variables (credit) and flow variables (GDP) in our analysis and this could somehow influence our results. In Appendix B, we show that the spectral densities for a stock and its flow should be the same, though there is a phase shift between the stock and the flow.

To cross-check the results from the multivariate spectral analysis, we use the Bry-Boschan quarterly algorithm for multiple time series to estimate the average duration of cycles in financial variables and real economic variables. This algorithm identifies the peaks and troughs in a time series given specific censoring parameters, such as the minimum length of a peak, expansion and contraction.[8] This method has the advantage of being able to estimate classical cycles rather than cycles in growth rates. However, it is not without its disadvantages; it is a purely statistical measure, with no underlying economic theory and it can be sensitive to the choice of censoring parameters, as shown in Mönch and Uhlig (2005).

If there is a financial cycle that is different from the macroeconomic cycle, as suggested in the literature, we would expect to find little overlap in the confidence intervals for the dominant frequencies in the multivariate spectral density for financial variables and economic variables, as well as large differences in the average duration of cycles in financial variables and economic variables as determined by the multivariate Bry-Boschan turning point algorithm.

Turning to how the financial and business cycles might be related, we first identify how often the financial cycle and the macroeconomic cycle are synchronised. For simplicity, we use credit growth to represent financial variables, and GDP growth to represent real economic variables. To obtain the cycles in these variables we use a

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7   For the bootstrapping, we use a Gaussian kernel to smooth the cross-spectra and uses a *Chi*-squared (2) distribution for the errors between the smoothed and non-smoothed cross-spectra. See Franke and Härdle (1992) and Berkowitz and Diebold (1998) for more information.

8   The conditions we use are consistent with those used in Harding and Pagan (2002); a cycle needs to be at least 5 quarters long, a contraction or expansion needs to be at least 2 quarters long and peaks and troughs need to alternate.

Christiano-Fitzgerald band-pass filter.[9] Two cycles in each time series are identified; a cycle between 8 and 30 years, often described as the range for financial cycle frequencies in the literature, and a cycle between 2½ and 8 years, often described as the range for macroeconomic cycles in the literature. We use the Bry-Boschan quarterly turning point algorithm on these cycle estimates to determine the degree of synchronisation between the cycles. If there is a financial (credit) cycle that is independent from the macroeconomic (GDP) cycle, or there is a significant phase shift between these cycles, we might expect the cycles to have a low degree of synchronisation.

Following this, we explore the lead/lag relationship between the cycles in credit growth and GDP growth. In particular, we employ cross-correlograms, cross-spectral densities and Granger causality tests to understand the relationship between these cycles. The results should help identify which cycle, if any, tends to lead the other and the number of lags involved. Given the discussion in the literature of a long financial cycle that leads to economic instability, credit growth could lead GDP growth at longer frequencies. At the same time, we expect that economic conditions matter for financial conditions and so GDP growth may also lead credit growth.
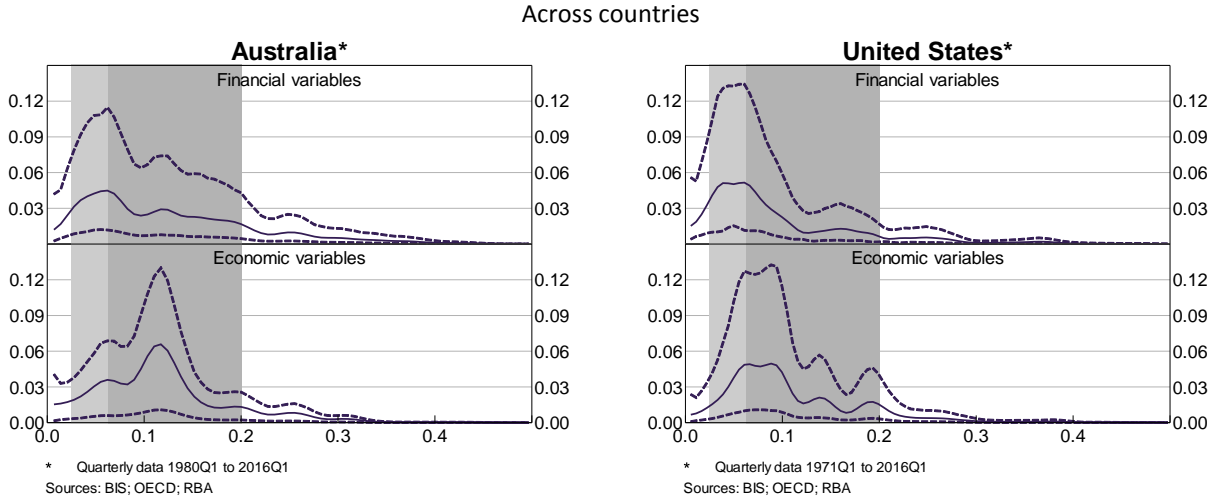
3.4 Results

3.4.1   Is there a financial cycle separate from the business cycle?

Figure 1 shows the measures of co-movement between the financial variables and the economic variables in the frequency domain for all five countries, while Table 1 presents the estimates of the dominant cycle lengths at which most of the co-movement occurs and the 95 per cent confidence intervals around these estimates. These estimates are used as an approximation for the length of the financial and business cycles.
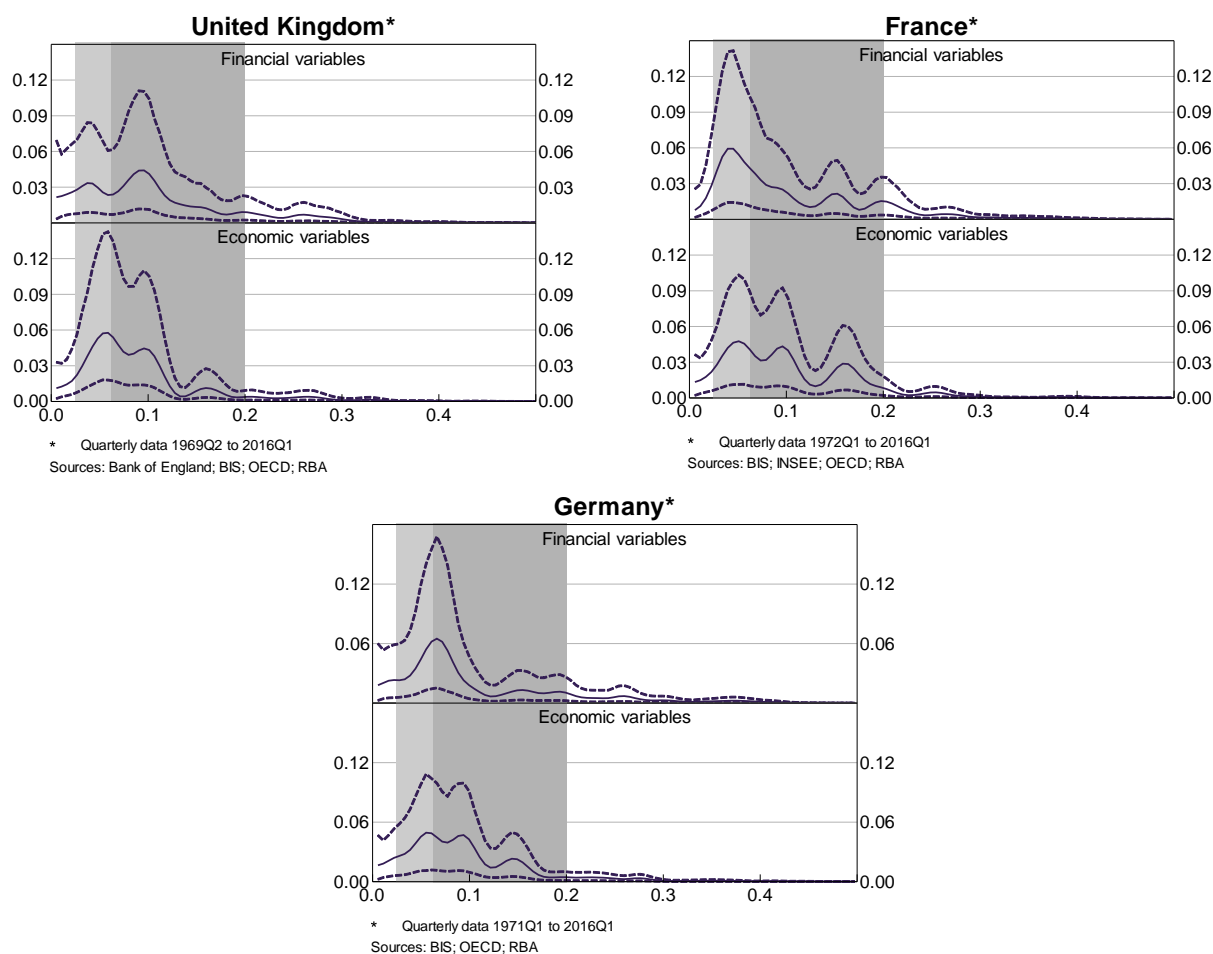
The results suggest that the common cycle in financial variables is different from the common cycle in economic variables, though this varies across countries. For the US, Australia and France, financial variables tend to co-move most at longer frequencies than economic variables; this is consistent with the literature that claims that financial cycles are longer than the business cycle, though the estimates are at the lower end of the range of financial cycle estimates in the literature. In contrast, financial variables in the UK and Germany tend to co-move most at shorter frequencies than economic variables. In several countries there is a secondary peak in the multivariate spectral density for economic variables that occurs at a shorter frequency, which could suggest the presence of shorter cycles in the data that are less dominant but nonetheless still important.[10]

Figure 1: Multivariate Spectral Densities

Across countries

*   Quarterly data 1980Q1 to 2016Q1
Sources: BIS; OECD; RBA

*   Quarterly data 1971Q1 to 2016Q1
Sources: BIS; OECD; RBA

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9   Similar results are found using a Hodrick-Prescott filter, where lambda is chosen based on optimising the unfiltered spectral densities between 8 and 30 years, and 2½ and 8 years. These filtering methods suffer from end-point problems, so care needs to be taken when interpreting the beginning and end of the filtered series.

10   However, the presence of two dominant cycles could also be the result of changes in the co-movement in the data over time.

United Kingdom*

* Quarterly data 1969Q2 to 2016Q1
Sources: Bank of England; BIS; OECD; RBA

France*

* Quarterly data 1972Q1 to 2016Q1
Sources: BIS; INSEE; OECD; RBA

Germany*

* Quarterly data 1971Q1 to 2016Q1
Sources: BIS; OECD; RBA

Notes: Smoothed with a Parzen window and standardised; dashed lines represent 95% confidence bands; darker (lighter) shaded area corresponds to cycle frequencies between 2½ and 8 years (8 and 30 years); financial variables are credit growth, housing price inflation, equity price growth, corporate bond spreads, change in 10 year government bond rates; economic variables are GDP growth, change in unemployment rate and employment growth

Table 1: Estimates of financial and business cycles

Peaks in multivariate spectral density

	Financial cycle estimate	Confidence interval	Business cycle estimate	Confidence interval
	Years	95%	Years	95%
US	8.2	[5.3, 22.6]	5.7	[4.5, 11.3]
Australia	8.1	[2.9, 18.1]	4.3	[3.6, 10.4]
UK	5.2	[4.3, 94.0]	8.5	[4.5, 15.7]
France	12.6	[3.5, 17.7]	9.8	[3.1, 14.8]
Germany	7.5	[5.7, 22.6]	9.1	[4.7, 18.1]

Sources: Bank of England; BIS; INSEE; OECD; RBA

As a comparison to the multivariate spectral analysis, we also estimate the average duration of cycles in financial and economic variables over the same time periods using the multivariate Bry-Boschan turning point algorithm. The results are presented in Table 2 and suggest that the average duration of the cycles in financial and economic variables has been broadly similar over the past few decades. In contrast to the results in Table 1, these results suggest that economic variables have generally had longer cycles than financial variables, particularly in Australia.

Table 2: Estimates of financial and business cycles

Average cycle lengths using Bry-Boschan quarterly algorithm

	Financial variables[a]			Economic variables[b]		
	Contraction	Expansion	Cycle	Contraction	Expansion	Cycle
US	2.1	3.7	**6.5**	1.2	4.7	**6.4**
Australia	3.1	2.5	**5.8**	1.8	7.4	**9.8**
UK	3.4	3.0	**6.4**	1.8	5.2	**7.0**
France	4.6	2.8	**7.3**	1.6	6.0	**8.5**
Germany	2.3	2.3	**4.6**	2.0	2.9	**5.1**

(a) Includes credit, housing prices, equity prices and 10-year government bond yields
(b) Includes GDP, employment and unemployment rate
Sources: Bank of England; BIS; INSEE; OECD; RBA

While the multivariate spectral densities are visually different, the large confidence intervals around the estimates for financial and business cycles and the considerable overlap between these confidence intervals cast uncertainty on the true cyclical behaviour of financial variables and economic variables. Given the overlap between the confidence intervals as well as the results from the turning point method, we cannot confidently conclude that the financial cycle is longer than the business cycle.

There are some weaknesses in our analysis that could affect the results. The first is that structural breaks may have occurred, leading to changes in the cycles over time; this is plausible given that our results cover a time period in which political and economic policy regimes changed significantly, including the reunification of Germany, financial liberalisation and changes in monetary regimes. The turning point method results should provide some robustness here, since the time-varying nature of this method implies that it would not be affected by structural breaks in the same way that the spectral analysis would be affected. As an additional robustness check, we repeat the multivariate spectral analysis using data from 1980 onwards for all countries (Table C2 in Appendix C).[11] While the change in samples does impact the results to some extent, particularly for the US and the UK, the confidence intervals around our estimates remain wide and still overlap.

Secondly, our choice of financial variables could influence our results. As documented in the literature, there appears to be consensus that credit growth and housing price inflation are more relevant when it comes to measuring the financial cycle than other financial variables, reflecting their superior ability to predict financial crises (Drehmann et al 2012; Jordà *et al* 2015). As a robustness check, we repeat our analysis excluding equity prices and long-term government bond rates. As shown in Table C2 and C3 in Appendix C, excluding these variables leads to an increase in the dominant frequencies in the multivariate spectral densities for financial variables as well as an increase in the average duration of cycles in financial variables, particularly in the US and UK. Nonetheless, there remains a significant overlap in the confidence intervals around the dominant frequencies in the multivariate spectral densities for financial and economic variables and the turning point method still estimates similar cycle lengths for some countries. Therefore, we believe that there is still uncertainty around the estimates of the financial and business cycles.

Finally, there is the issue of insufficient data. If, as claimed by other papers, there exists a longer financial cycle and we have data covering less than five decades, there will only be a limited number of cycles in the data and it will be difficult to confidently identify the length of the cycle. As another robustness check, we repeat the multivariate spectral analysis using the long-run annual database provided by Jordà, Schularick and Taylor (forthcoming).[12,13]

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11 We choose 1980 for a number of reasons: a long enough sample is required for stationarity purposes and in order to capture longer cycles; and around this period many countries went through a period of financial liberalisation, so there likely to be a structural break between the 1970s and the 1980s.

12 This database covers the period 1870 to 2013. Due to data availability, the time period varies for each country. Note that the only economic variable used in this robustness check is GDP growth, since this database does not provide employment or unemployment data.

The results are provided in Table C4 of Appendix C. These results are not entirely consistent with the quarterly data, particularly for the UK and Australia, which is unsurprising given that the annual data cover a much longer time period than the quarterly data and many structural breaks are likely to have occurred over this period. Notably, the confidence intervals remain large around the dominant frequencies for financial variables and economic variables, implying that, even with a longer dataset, we cannot verify that the financial variables have different cyclical behaviour than economic variables.

Given that the robustness checks do not provide evidence that cycles in financial variables vary considerably from cycles in economic variables, we conclude that our results do not provide sufficient evidence for the existence of a financial cycle that is longer than the business cycle. This is not to say that this claim cannot be true, but rather that we cannot find enough evidence to support this claim.

### 3.4.2   How are the financial variables and business variables related?

Despite the fact that we do not have evidence that the cyclical behaviour of financial variables and economic variables differs considerably, it is still important to understand the relationship between these variables. It is possible that these variables do have different cyclical behaviour but we cannot prove it due to a lack of data, meaning that there would undoubtedly be times were their cycles are in different phases of the cycle; and, even if their cyclical properties are not different, these cycles could be in different phases at the same time. In fact, just by observing current and recent events across many countries suggests that developments in financial variables can be at odds with developments in economic variables. If there is a relationship between these cycles, knowing how often the cycles are in the same phase will be useful to understand how often there might be conflict between the cycles (and the policies that manage them).

Figure 2 presents estimated cycles in credit growth and GDP growth across the five countries, using the Christiano-Fitzgerald band-pass filter; the longer and shorter cycles shown in these graphs correspond to cycles of 8 to 30 years and 2½ to 8 years, respectively. It is clear that the cycles in credit growth and GDP growth are similar, though they are sometimes out of sync. The longer cycles in credit growth and GDP growth appear to be quite synchronised in the US and Australia, and it looks as though these cycles are related.

Figure 2: Cycles in credit growth and GDP growth

Across countries, quarterly data, year-ended

Sources: BIS; OECD; RBA

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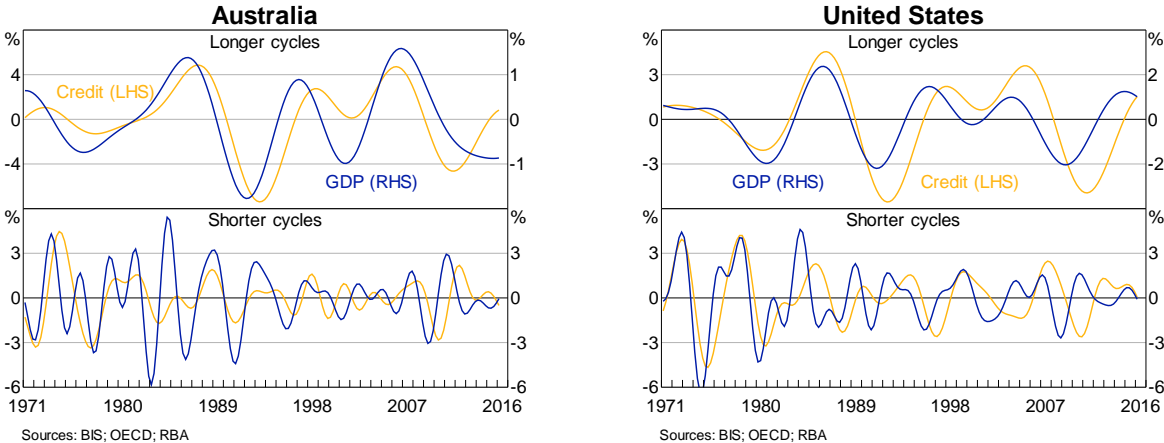
13 The turning point method is typically used on monthly and quarterly data; annual data will only be able to picks up major contractions in the cycles, and therefore the cycles are typically longer (Harding 2002). As such, we do not run the robustness check using annual data using the turning point method.

United Kingdom

France

Germany

Sources: BIS; OECD; RBA
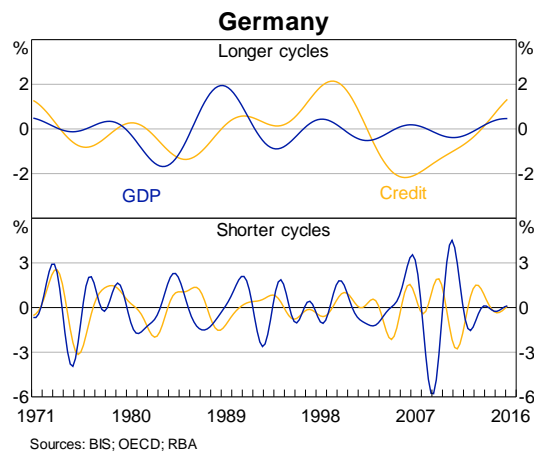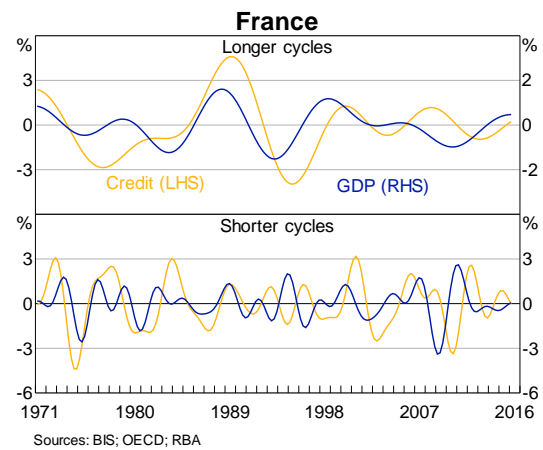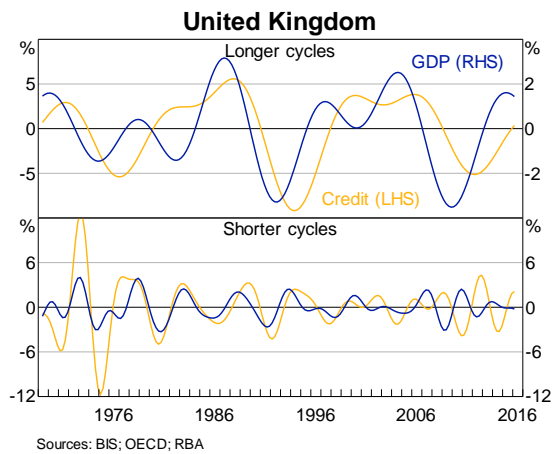
Notes: Cycles are obtained using the Christiano-Fitzgerald band-pass filter; longer cycles correspond to cycle frequencies between 8 and 30 years; shorter cycles correspond to cycle frequencies between 2½ and 8 years

Using the Bry-Boschan turning point algorithm updated in Harding and Pagan (2002), Table 4 presents the average duration of contractions and expansions in longer cycles and shorter cycles in credit growth and GDP growth as well as the degree of synchronisation between these cycles. The length of expansions and contractions are similar for the cycles in credit growth and GDP growth, suggesting a similar length in the cycles.

In terms of synchronisation, the longer cycles in GDP growth and credit growth have a higher degree of synchronisation than the shorter cycles. Compared with two random series, which you would expect to be in sync 50 per cent of the time, the longer cycles were in sync 67 per cent of the time on average, while the shorter cycles were in sync 57 per cent of the time.[14] For the US and Australia, the longer cycles are in the same phase for around three-quarters of the time. The high degree of synchronisation between the cycles in the US and Australia could reflect the importance of the financial services sector for its economy, though it is interesting that we do not find a similar result for the UK, which also has a large financial sector. In addition to looking at synchronisation *within* countries, we can also estimate the degree of synchronisation *across* countries. Table D1 in Appendix D presents these results and shows that the cycles in some countries can be reasonably synchronised.

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14 We would expect the concordance statistic to be 0.5 for two random series since they do not have a unit root (i.e. there is no drift in the series). It takes drift common to both series to increase the expectation of the concordance statistic above 0.5 (McDermott and Scott 2000).

Table 3: Synchronisation between cycles

Average cycle lengths and concordance statistics using Bry-Boschan quarterly algorithm

	Credit growth		GDP growth		Degree of synchronisation
	Contraction	Expansion	Contraction	Expansion	Concordance statistic
Longer cycles					
Australia	4.8	6.1	5.2	5.2	0.73
US	5.8	4.4	4.4	4.3	0.77
UK	5.0	4.2	5.0	4.7	0.57
France	4.3	4.7	4.7	4.3	0.65
Germany	5.1	4.4	6.2	4.6	0.64
Average	**5.0**	**4.8**	**5.1**	**4.6**	**0.67**
Shorter cycles					
Australia	1.6	1.5	1.8	1.5	0.61
US	2.3	1.7	2.2	1.8	0.55
UK	2.3	1.9	2.1	1.9	0.60
France	1.7	1.5	1.8	1.6	0.51
Germany	1.9	1.9	2.1	2.1	0.54
Average	**2.0**	**1.7**	**2.0**	**1.8**	**0.57**

Notes: Longer cycles correspond to cycle frequencies between 8 and 30 years; shorter cycles correspond to cycle frequencies between 2½ and 8 years; concordance statistics represent the proportion of time the cycles are in the same phase
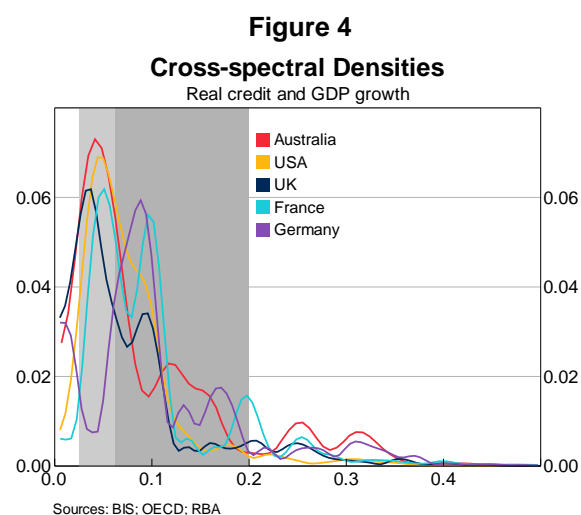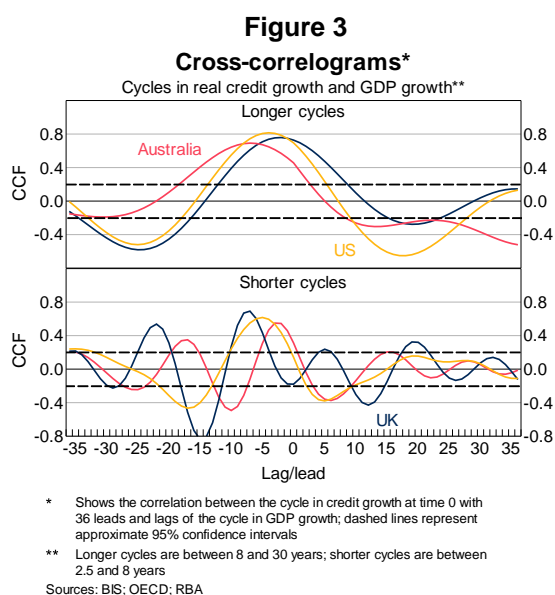
Sources: BIS; OECD; RBA

While the cycles in credit and GDP growth have often been in the same phase, there have still been times when they are in different phases. What this means for policymakers depends on the relationship between the cycles. The literature discussed in Sections 2 and 3.1 provides many reasons as to why a relationship may exist between developments in the financial system and the economy. If we do find evidence of a relationship, developing a better understanding of this relationship should help policymakers to forecast the future paths of the cycles in credit and GDP (and, by extension, the financial and business cycles) and set policy accordingly.

A visual inspection of the time series estimates of the cycles suggests that cycles in credit growth and cycles in GDP growth are related. This can also be seen in the cross-correlation between the cycles. In general, we find that lags in the cycles of GDP growth are positively correlated with the cycles in credit growth up to a few years ahead, particularly in the longer cycles (Figure 3).[15] This suggests that cycles in GDP growth could positively lead cycles in credit growth. At the same time, some longer leads/lags in GDP growth cycles are negatively correlated with credit growth cycles. This suggests that cycles in credit growth could negatively lead cycles in GDP growth. Therefore, the relationship could run both ways or there could just be a one-way relationship with a phase shift.

Moving from the time domain to the frequency domain, Figure 4 presents cross-spectral densities for GDP growth and credit growth, which are effectively frequency decompositions of the cross-correlation between the cycles. These graphs suggest that the co-movement between GDP growth and credit growth is strongest at lower frequencies; that is, GDP growth and credit growth tend to co-move more in the longer run. The exception here is Germany, where a significant amount of cross-correlation occurs at frequencies between 2½ and 8 years. Using the phase spectrum, we can determine the lead/lag relationship between GDP growth and credit growth at different frequencies.[16] While the phase spectrum can often be difficult to interpret, our results suggest that credit growth tends to lead GDP growth at longer frequencies.

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15 For the cross-correlograms for France and Germany, see Figure D1 in Appendix D.

16 The phase spectrum can only be meaningfully interpreted at frequencies for which spectral mass is high.

Figure 3
Cross-correlograms*
Cycles in real credit growth and GDP growth**

Figure 4
Cross-spectral Densities
Real credit and GDP growth

* Shows the correlation between the cycle in credit growth at time 0 with 36 leads and lags of the cycle in GDP growth; dashed lines represent approximate 95% confidence intervals

** Longer cycles are between 8 and 30 years; shorter cycles are between 2.5 and 8 years

Sources: BIS; OECD; RBA

Sources: BIS; OECD; RBA

So far, we have only looked at the correlation between the cycles; this does not prove (nor rule out the possibility of) a relationship between credit growth and GDP growth. Assuming a relationship does exist and abstracting from other variables relevant in this relationship, there are two plausible cases of a causal relationship occurring between credit growth and GDP growth that are consistent with our results and the theory:

- **GDP growth → Credit growth**. Changes in GDP growth are independent from changes in credit growth and cause changes in credit growth. The ability of GDP growth to affect credit growth is intuitive in terms of economic theory and suggested in the results above. In this case, GDP growth leads credit growth and the negative correlation between credit growth and the lead/lags of GDP growth merely reflects GDP growth leading credit growth with a phase shift.

- **Credit growth ↔ GDP growth**. The theory, cross-correlogram results and time series estimates suggest that GDP growth can lead credit growth. However, as discussed earlier in the paper it is also entirely reasonable to expect that credit growth can influence GDP growth and this could help explain some of the results from the cross-correlograms and cross-spectral densities. There is potential for both a positive and negative relationship to exist. For example, a strong rise in credit growth may be associated with a build-up in risks that, if they materialise, could have significant effects for the financial system and the economy. There is also the potential for positive effects for the economy during these expansions, including the extension of credit to productive and innovative ventures.

To formally test whether the relationship is significantly one-way or whether it can run both ways, we employ Granger Causality tests. Table 6 presents the *p*-values from the tests for whether GDP growth/credit growth Granger Causes credit growth/GDP growth.[17] Overall, the results suggest that cycles in GDP growth Granger cause cycles credit growth in all countries. For some countries there is evidence that cycles in credit growth and GDP growth Granger Cause each other, suggesting that the relationship could go both ways.

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17 This analysis only includes data from 1980 onwards, since we believe that the nature of the relationship between GDP growth and credit growth is likely to have changed during financial liberalisation

Table 4: Granger Causality Tests

p-values, 1980:1 to 2016:1

Alternative hypothesis	Australia	US	UK	France	Germany
Actual data					
Credit growth → GDP growth	0.04	0.14	0.75	0.06	0.29
GDP growth → credit growth	0.03	0.03	0.01	0.04	0.13
Longer cycles[a]					
Credit growth → GDP growth	0.06	0.05	0.39	0.24	0.01
GDP growth → credit growth	<0.00	<0.00	<0.00	<0.00	<0.00
Shorter cycles[a]					
Credit growth → GDP growth	0.01	0.18	0.23	0.58	0.09
GDP growth → credit growth	0.10	0.06	<0.00	<0.00	0.08

(a) Since Granger causality tests are not possible with band-pass filters, these cycles are identified using an HP filter with lambda chosen by optimising the spectral density of the actual series between 2½ and 8 years for shorter cycles and 8 and 30 years for longer cycles.
Sources: BIS; OECD; RBA

Overall, it is clear that there is some lead/lag relationship between credit growth and GDP growth. The results clearly suggest that GDP growth leads credit growth, but at the same time they do not rule out the possibility that credit growth leads GDP growth. The mixed results imply that we cannot confirm whether the financial cycle leads or lags the macroeconomic cycle in long run or short run. In a way, the mixed results are not surprising given the discussion on the links between the financial system and the economy at the beginning of this paper; there are explanations for the financial system affecting the macroeconomy and vice versa. It is likely that the relationship between GDP growth and credit growth depends on the circumstances and other factors not captured in this analysis, meaning that in some cases GDP growth will lead credit growth and in other cases credit growth will lead GDP growth. Furthermore, the interpretation of these results is further complicated by the potential for regime switching as well as the presence of the financial liberalisation in the data. However, what is clear from our results is that the cycles are unlikely to be independent of each other and will sometimes be moving in different directions.

4. Policy Implications

The analysis presented in this paper suggests that there is little evidence to conclude that cycles in financial and activity variables operate at different frequencies – in particular, it is difficult to argue that financial variables have a cycle that is longer than the cycle for economic activity variables. It is worth noting, however, that we cannot definitively conclude that the two cycles have the same frequency either and the results may be subject to type II error (failure to reject the hypothesis that the peak frequencies are the same when they are, in fact, different). Furthermore, the evidence suggests that developments in activity variables lead those in financial variables, but it is more difficult to conclude the converse. While the power of these tests may not be high, the results suggest that much more evidence is required to establish whether the cycles in these variables are distinct from one another, and if so, why that might be the case.

We have chosen to focus our analysis on the differences in the frequencies and phases in the financial and business cycles rather than the amplitudes in these cycles. However, we acknowledge that there could be policy arguments for being concerned about the amplitudes of the cycles, given that the financial accelerator mechanism implies that larger amplitudes in the financial cycle may result in larger fluctuations in the business cycle (and vice versa). On the other hand, it is unclear what the welfare costs of a credit cycle with higher amplitude might be if it means that households are using credit to smooth consumption.

Overall, for many of the economies considered in our analysis, the cycles in activity and financial variables appear to fluctuate at the same frequency. However, the analysis does not give us any insights into what drives these

cycles and the mechanisms through which the two cycles affect each other. Unfortunately, our analysis, and those of many other authors, has not been able to identify the drivers of these cycles, so we are limited to consider general scenarios.

Given the effect economic activity has on financial variables and vice versa, it is highly unlikely that the two cycles are independent of each other and have similar frequencies by chance. It is very likely that if the two cycles had a similar frequency, they would have common drivers. As a result, financial variables should be useful in understanding developments in the business cycle.

If the main drivers of the two cycles were real (e.g. productivity), then fluctuations in financial variables, particularly credit, would reflect fluctuations in the need to finance economic activity. Financial variables would be used by policymakers to make an inference about the state of the economy. In this view of the world, the standard policy tools that have typically been used to address cyclical fluctuations in activity, namely fiscal and monetary policy, would continue to be used to 'smooth' out the business cycle and minimise variations in output growth and unemployment, and minimise deviations of inflation from its target. The extent to which fluctuations in each are tolerated will depend on the weights policymakers place on each. Policymakers would only be concerned with and respond to fluctuations in financial variables to the extent that these variables are informative about economic conditions.

However, it is possible that not all of the drivers of these cycles are real and some could be financial in nature; in this case, there could be grounds for policymakers to address both real and financial drivers. The key question is what these drivers could be. In this paper, we do not discuss changes to prudential regulation and financial system deregulation, which have affected both real and financial variables and may distort the frequencies that we detect in the data, particularly given the relatively short period for which data are available. That being said, these financial drivers can be difficult to identify, particularly those that are independent of the state of the business cycle.

Many of the financial drivers identified in the literature relate to sentiment, particularly regarding risk, and expectations (particularly for asset prices). One mechanism by which a financial cycle can affect the business cycle is in changes in the perception of risk that can affect the propensity for agents to underestimate the risks they are taking during economic expansions. This leads to agents taking on more risk than they would otherwise. This risk taking is not just reflected in the type of investments or portfolio allocations they make, but also in the amount of leverage they are willing to use to fund these activities. This can be reflected in 'overvalued' assets, expectations of sustained higher asset price inflation, unsustainable levels of debt and lower credit spreads. This situation can ultimately, but not always, lead to higher loan defaults and a contraction in bank lending that ultimately results in a contraction of economic activity. This is not to say that it is just households and firms that take these risks; the financial sector can also underestimate risks that lead to imprudent management of its balance sheet. Once growth slows, or even a recession begins, perceptions of risk adjust and could even lead to agents overestimating risk leading to higher spreads and expectations of asset price deflation that can exacerbate the amplitude of the business cycle. The main question is: what causes these changes in perceptions in risk and expectations of asset price inflation? Real shocks to the economy will play an important role.

Some have argued that the conduct of monetary policy can contribute to the cycles we observe in financial variables (Borio 2012, Juselius *et al* 2016). Borio (2012) argues that monetary policy that is too narrowly focussed on inflation may not necessarily tighten policy when financial 'booms' occur because inflation is low and stable. In a scenario when the potential growth rate of the economy rises, credit growth and asset price inflation rise, but the spare capacity in the economy would constrain the monetary authorities' ability to respond. We do not find this scenario convincing. The cause of the increase in the supply potential of the economy matters. If the potential growth rate of the economy increased due to a rise in productivity growth, this would raise the neutral real interest rate in the economy and soon enough, the monetary authority will realise the need to begin to tighten monetary policy, which would reign in any boom in credit and asset prices.

Although the business and financial cycles appear to be closely linked and operate at similar frequencies, it is possible that they will be in different phases. In particular, there will be times when the economy is slowing (including the case where growth is positive but below potential), but there is reasonably high asset price inflation and accelerating credit. This raises the question of how deal with such situations, whether it should be dealt with at all, and if so, what are the appropriate policy tools to address this confluence of events. The main research question, which we do not address here, is: what are the market distortions and potential market failures that lead to this, admittedly infrequent, confluence of events? Why would agents continue to borrow (presumably unsustainably)? Why would the financial system continue to lend in this environment? Why are some asset prices not responding to slower growth in the economy? In a first-best world where we understand the answers to these questions and unconstrained, we would be able to design policies to address the distortions and market failures that produce these outcomes. However, we choose to discuss the role of monetary policy in addressing 'financial imbalances' given the suggestions that monetary policy should be conducted to moderate fluctuations in financial variables.

4.1 Monetary policy and the financial cycle

Monetary policy has traditionally been used to deal with the business cycles. It is particularly good at doing this because interest rates provide the price signal for agents to intertemporally reallocate resources; during booms, interest rates rise and encourage agents to push economic activity into the future, while during downturns, it encourages agents to bring economic activity forward from the future. It does this to both keep unemployment around the non-accelerating inflation rate of unemployment (NAIRU) and inflation around its target. However, by adjusting interest rates, monetary policy also has an effect on credit growth and asset price inflation. So if the economic conditions and inflation warrant lower interest rates but credit growth and asset price inflation are still 'too high', should monetary policy be used to curtail the potential risks this might be posed to economic activity in the future?

Traditionally, the policy making framework has focussed on the role of monetary and fiscal policy to moderate the business cycle and prudential policy to address the risks in financial institutions.[18] The risks discussed in the financial cycle literature do not just consider the risks on financial institutions' balance sheets, but also the risks on private sector balance sheets. These risks can be magnified by asset prices that are too high and inflating at a rapid rate and credit growing at an 'unsustainable' pace.

A laissez-faire view of this problem is that markets are efficient and that households and firms know the risks they are taking; they will bear the consequences if those risks materialise. Therefore, to the extent that these risks are idiosyncratic and uncorrelated, and that prudential policy has been applied appropriately, there is no need for further policy intervention. Another argument, however, is that these risks across households and firms are correlated and the externalities of a number of households and firms experiencing financial distress at once are large and negative – those who were more prudent in their decision making would still be affected if there were an economic downturn. Under such a scenario, there should be a policy intervention to ensure that these particular risks are managed appropriately and the consequences of such a scenario are minimised.

The conundrum for monetary policy is that there are a number of channels by which it can affect risk taking behaviour. Lower interest rates can incentivise banks to take on more risks such as higher leverage, greater maturity mismatch and to ease lending standards (Gambacorta 2009, Agur and Demertzis 2015). Lower interest rates can also encourage investors to 'search for yield' and to expose themselves to more risk (Rajan 2005). Together, these are often referred to as the 'risk-taking channel' of monetary policy (Borio and Zhu 2012, Bruno and Shin 2015). In addition, higher asset prices resulting from lower rates increases the value of collateral, potentially encouraging greater borrowing and creating an asset price boom fuelled by credit. Alternatively, higher interest rates increase debt repayment burdens, which can potentially lead to higher default rates (Sengupta 2010). Higher interest rates also put pressure on banks' funding costs, thereby compressing their profit margins

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18 We will mostly ignore the role of fiscal policy in this paper.

and potentially encouraging them to engage in riskier lending (Agur and Demertzis 2015). Exchange rate appreciations associated with interest rate increases can encourage foreign currency borrowing, which can pose significant risks to unhedged foreign-currency borrowers during any subsequent depreciation.

Agur and Demertzis (2015) finds that the 'risk-taking channel' tends to dominate banks' risk-taking behaviour during the upswing of a financial cycle, while the profitability channel tends to dominate during the contraction phase. Drawing on this result, the scenario outlined above where the macroeconomy is in a contractionary phase but the financial cycle is in an expansionary phase would put some pressure on the central bank to ease policy, but this could encourage more risk taking at a time when these risks may already be unacceptably high. The choice a central bank makes in this situation will depend on how much value it places on achieving its objectives in the long-term versus the short-term. In the case of an inflation-targeting central bank, the other consideration would be the effect on inflation expectations if a deviation from target was protracted enough to shift expectations and make the inflation objective harder to achieve in the long-term without some economic cost.

Despite the potential for short-term conflict, in the long run price stability should have some positive spillovers for financial stability and vice versa, since a stable macroeconomic environment and a stable financial environment are necessary conditions for each other (Smets 2013). Furthermore, to the extent that easier monetary policy can raise inflation and nominal income growth, this should help reduce the debt repayment burdens for borrowers. Moreover, during a downturn, some risk taking may be desirable to encourage more investment and growth. The difficulty for central banks is determining what amount of risk is desirable when those financial risks become material enough to potentially pose a threat to macroeconomic stability. This is where prudential policy plays an important role. Smets (2013) and Maddaloni and Peydro (2012) find evidence that the effect of lowering interest rates on banks' risk-taking is reduced under more rigorous prudential policy. Therefore, one avenue to manage these risks is through prudential policy.

This begs the question as to what instruments should be used to address price and economic stability, and can we rely on one instrument to manage the business and financial cycles (assuming there are times they are in conflict). Tinbergen (1952) suggests that the number of policy instruments that should be used is the same as the number of policy objectives. Policy instruments should be used to achieve the objectives they can most efficiently attain. This principle would suggest that monetary policy should focus on targeting price stability, where it has been proven effective, and leave other instruments to target other policy objectives. Since the nominal interest rate is only one tool, can it be expected to effectively achieve its traditional objectives in addition to managing the financial cycle?

4.2 Macroprudential policy

Reflecting its blunt nature, many papers have concluded that using monetary policy to target concerns regarding risks to the financial system would require large changes in interest rates, which would be costly and inefficient, particularly when the risks are building up in a specific area of the financial system (Alpanda, Cateau and Meh 2014, Svensson 2016). In contrast, macroprudential policy consists of a range of tools, which means macroprudential policy can have greater flexibility to manage the financial cycle and allow monetary policy to focus on its primary objectives. The most common macroprudential tools include caps on loan-to-valuation ratios (LVRs), minimum serviceability requirements and capital requirements. These tools can be targeted to address the build-up of risks for particular sectors – such as the household sector – and strengthen the financial system's resilience.

It is unlikely that macroprudential policy is capable of perfectly addressing all types of risks that build during the expansionary phase of a financial cycle. One clear side-effect of using macroprudential policy to restrict lending to a particular sector of the economy is the diversion of lending to unregulated financial institutions (the shadow banking system) where policymakers tend to have limited influence; in this case, a blunt tool such as monetary policy could be more effective by controlling the cost of funds (according to Stein 2013, monetary policy 'gets in all the cracks').

Another drawback is the uncertainty around the effectiveness of some macroprudential policies. There is currently mixed evidence of the effectiveness of macroprudential tools, particularly in advanced economies, and effectiveness is likely to vary depending on the type of tool used and the circumstances in which it is used (Cerutti *et al* 2015). Moreover, it can be difficult to measure the effectiveness of macroprudential tools for a number of reasons, including the lack a clear and quantifiable target and limited experience with using these tools in the post-crisis period. Furthermore, it is not entirely clear under what circumstances these instruments would be more effective and when they would be ineffective. For example, Guibourg *et al* (2015) suggests that countercyclical capital requirements would not have been sufficient to stop rapid credit growth in the lead up to the financial crisis in the UK and Spain.

There are significant institutional issues related to macroprudential policy. For the macroprudential tools to be a useful part of the policy arsenal, it must be clear that they can be used without political interference – i.e. that the responsible authority has the independence to use the tools at its disposal. There is the question of who should be in control of these instruments, particularly when the financial regulator and central bank are separate. Macroprudential policy determined by the central bank would require the financial regulator to implement these policies. It would be unclear, in this instance, who would be accountable for the use of the tools. Finally, there is a question as to how quickly some policies can be implemented and enforced – if the time required is too long, risks may have materialised before the policies could be implemented, by which time the policies may need to be abandoned.

The other drawback of trying to use macroprudential tools to manage financial risks in the economy is the risk profile of projects being undertaken as lending to the private sector becomes a binding constraint on economic activity. While these policies may be effective, they can be just as blunt as monetary policy. This means that some viable, albeit riskier, investments may be denied funding, which could result in less funding for innovation that would have otherwise helped to drive higher productivity growth.

4.3 Implications for monetary policy

Despite these potential limitations and costs, macroprudential policy would still be a valuable complement by helping to manage some risks as monetary policy tries to stabilise economic activity. However, this means that we need to consider how monetary and macroprudential policies might interact. These policies should move together and reinforce each other when the cycles are in sync. However, when the business and financial cycles are out of sync, these policies may pull in different directions and counteract the effect of the other. We know from the previous section that the stance of monetary policy can influence the build-up of risks and therefore the appropriate stance of macroprudential policy. What is of interest in this section is how changes in macropudential policy, responding to a change in financial imbalances either related or unrelated to the stance of monetary policy, will affect the transmission of monetary policy.

To some extent, this type of discussion is not new; there has been extensive research on the interactions between policies, most notably monetary and fiscal policy (the most obvious example being Sargent and Wallace 1981). In the case of monetary and fiscal policy, both policies arguably target the business cycle and therefore have similar objectives; however, these policies have different tools and any conflict between these policies tends to come from political constraints.

We expect there to be some overlap in transmission channels between monetary and macroprudential policies since both policies work to achieve their objectives through affecting the price and quantity of credit in the economy, and through asset prices. The extent to which the interactions between these policies are significant will depend on the nature of the macroprudential tool that is being used (Dunstan 2014). In particular, it seems reasonable to expect that a broad-based macroprudential tool will have different implications for monetary policy than a narrow tool (i.e. targets specific sector, institution or market).

Macroprudential tools that affect the distribution of credit, such as LVR caps, will likely affect the functioning of monetary policy's wealth and balance sheet channels since these tools are most effective in tempering rapid credit

and asset price growth. For example, a tightening in these tools can reduce the effect of lower interest rates on housing price inflation, which can lead to lower household consumption through the wealth effect and tighten borrowing constraints through lower collateral values. As a result, the demand and inflation outlook are likely to be lower than if a tightening of macroprudential policy had not occurred.

Capital-based macroprudential tools can alter the price and supply of credit since equity funding is generally more expensive than debt funding (Harimohan and Nelson 2014). In response to higher capital requirements, banks have the option to increase retained earnings through lower dividend payouts, issue new capital and pass through higher funding costs to borrowers in the form of higher lending rates, or change the composition and/or size of their lending portfolio. All options can have implications for the outlook for demand and inflation; higher lending rates offset some of the effects of accommodative monetary policy to encourage borrowing and current consumption, while changes to banks' assets could reduce the supply of credit. If banks do choose to raise new capital, it is possible that more intense competition can limit their ability to pass on higher lending rates to borrowers.

Liquidity-based tools can affect banks' holding of illiquid assets, altering credit supply (e.g. long-term lending) and, to the extent they weigh on bank profitability, may result in higher lending rates. In addition, banks may have to change their funding profiles which can make funding more expensive, which would also be passed on to customers (through lower deposit rates and higher lending rates). This would have the effect of reducing the effectiveness of monetary policy.

When the cycles are in sync, marcoprudential policy should reinforce monetary policy. When the cycles are out of sync, macroprudential policy can work against monetary policy through counteracting its effects on the price and quantity of credit in the economy, and therefore domestic demand and inflation. However, this will vary depending on the type and scope of the macroprudential tools used, and some coordination of monetary and macroprudential policies will be required to ensure that all objectives can be achieved and all trade-offs are managed over an appropriate horizon.

Despite the drawbacks of macroprudential policy as well as the potential for monetary and macroprudential to interact, we believe that it is not ideal to use monetary policy to manage the financial cycle by itself; it seems reasonable to think that macroprudential policy should play a part. There are, however, some reasonable arguments for leaving the door open for central banks to consider its role in the build-up of financial risks when considering the appropriate course of action for monetary policy. However, we would agree with the view expressed by Peter Praet (2016) that the first line of defence must be a strong institutional and legal framework that directly targets the sources of market failure.

5. Conclusion

Since the financial crisis, there has been substantial work to identify the causes of the crisis and understand how future crises can be prevented. Among this work, it has been proposed, and has become commonly accepted, that there is a financial cycle that is longer than the business cycle and should be managed separately from the business cycle. Our results, on the other hand, suggest that there is not enough evidence to suggest that the financial cycle is longer than the business cycle. Nonetheless, even if these cycles are similar in length, there could still be reason for policymakers to address developments in the financial system. We find that the financial and business cycles are not always synchronised and can influence each other. Generally, we expect the business cycle will lead the financial cycle, but our results and observations from financial crises suggest that it is possible for financial developments to have a significant effect on economic activity.

We agree that monetary policy does have an effect on the 'financial cycle' but argue that there are limitations in its ability to manage the business and financial cycles at the same time. Although it is not a perfect solution, macroprudential policy has the flexibility to help manage the financial cycle, but if it were to be used, it would need to be coordinated with monetary policy to achieving the economic and financial objectives, and that the trade-offs are being managed appropriately across all these instruments.

What our analysis does not do, nor do many other papers, is to explain the drivers of these cycles. It is worth saying at the outset, that a very lengthy research agenda in the profession has tried to explain business cycles, and while our understanding has improved in this process, our knowledge is still fairly limited. If that experience is anything to go by, it will take as a long time to get a decent understanding of what drives financial cycles, if they exist.

Appendix A: Data sources

Table 1: Data sources and available data

	United States	Australia	United Kingdom	France	Germany
Credit	BIS	BIS	BIS	BIS	BIS
	(1952:4 – 2016:1)	(1953:4 – 2016:1)	(1963:1 – 2016:1)	(1969:4 – 2016:1)	(1948:4 – 2016:1)
Housing prices	BIS	BIS	OECD	BIS	BIS
	(1970:1 – 2016:1)	(1970:1 – 2016:1)	(1968:2 – 2016:1)	(1970:1 – 2016:1)	(1970:1 – 2016:1)
Equity prices	OECD	OECD	OECD	OECD	OECD
	(1957:1 – 2016:1)	(1958:1 – 2016:1)	(1958:1 – 2016:1)	(1957:1 – 2016:1)	(1960:1 – 2016:1)
10-year government bond yield	OECD	OECD	OECD	OECD	OECD
	(1953:2 – 2016:1)	(1969:3 – 2016:1)	(1960:1 – 2016:1)	(1960:1 – 2016:1)	(1956:2 – 2016:1)
GDP	OECD	OECD	OECD	OECD	OECD
	(1947:1 – 2016:1)	(1959:3 – 2016:1)	(1955:1 – 2016:1)	(1949:1 – 2016:1)	(1970:1 – 2016:1)
Employment	OECD	OECD	Bank of England	INSEE	OECD
	(1955:1 – 2016:1)	(1978:2 – 2016:1)	(1950:3 – 2016:1)	(1970:4 – 2016:1)	(1962:1 – 2016:1)
Unemployment rate	OECD	OECD	OECD	OECD	OECD
	(1955:1 – 2016:1)	(1966:3 – 2016:1)	(1955:1 – 2016:1)	(1967:4 – 2016:1)	(1968:4 – 2016:1)
CPI	OECD	OECD	OECD	OECD	OECD
	(1955:1 – 2016:1)	(1948:4 – 2016:1)	(1955:1 – 2016:1)	(1955:1 – 2016:1)	(1955:1 – 2016:1)

Source: RBA

Table A5: Summary Statistics
1970:Q1 - 2016:Q1[a]

	US	UK	Australia	France	Germany
Financial variables					
Credit growth	3.4	4.6	5.7	3.7	2.7
	(3.4)	(6.2)	(3.6)	(2.9)	(2.8)
House price growth	1.6	3.3	2.8	1.9	-0.3
	(6.1)	(9.4)	(6.4)	(4.8)	(2.5)
Equity price growth	2.2	1.1	0.3	2.2	2.0
	(16.4)	(19.2)	(19.6)	(22.3)	(20.6)
10-year Government bond yield (Δ)	-0.1	-0.1	-0.1	-0.1	-0.1
	(1.2)	(1.4)	(1.3)	(1.2)	(1.0)
Economic variables					
GDP growth	2.2	2.4	3.1	2.1	1.7
	(2.5)	(2.9)	(2.7)	(2.0)	(2.4)
Employment growth	0.1	0.3	0.3	0.6	0.4
	(1.4)	(1.4)	(1.5)	(1.4)	(1.7)
Unemployment rate (Δ)	0.0	0.1	0.1	0.2	0.1
	(1.1)	(1.0)	(1.0)	(0.7)	(0.9)

Notes   This table presents the means and standard deviation (in parentheses); all nominal variables are expressed in real terms using the CPI

(a)   Some variables cover shorter time periods. The house price growth data starts in 1971:1, with an exception for the UK. All data for France starts in 1971:1, except for the changes in the 10-year government bond yield and unemployment rate. The employment growth data are only available from 1978:1 and 1979:1 for the US and Australia, respectively.

Sources: Bank of England; BIS; INSEE; OECD; RBA
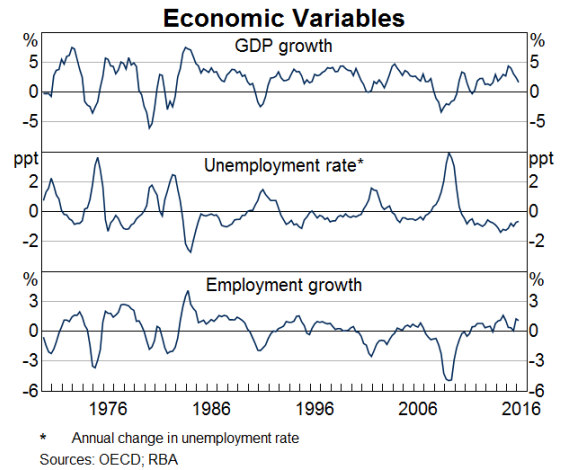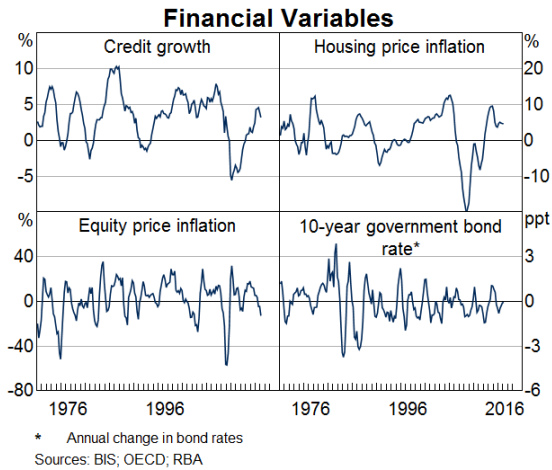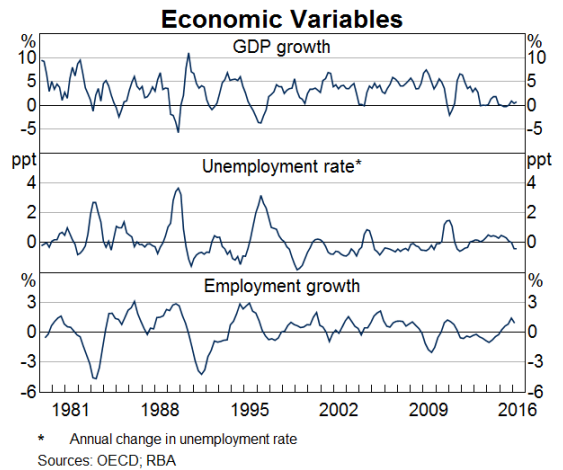
Figure A1: United States data

Financial Variables

Credit growth | Housing price inflation
Equity price inflation | 10-year government bond rate*

* Annual change in bond rates
Sources: BIS; OECD; RBA

Economic Variables

GDP growth
Unemployment rate*
Employment growth

* Annual change in unemployment rate
Sources: OECD; RBA

Figure A2: Australia data

Financial Variables

Credit growth | Housing price inflation
Equity price inflation | 10-year government bond rate*

* Annual change in bond rates
Sources: BIS; OECD; RBA

Economic Variables

GDP growth
Unemployment rate*
Employment growth

* Annual change in unemployment rate
Sources: OECD; RBA

Figure A3: UK data

Financial Variables

Credit growth | Housing price inflation
Equity price inflation | 10-year government bond rate*

* Annual change in bond rates
Sources: BIS; OECD; RBA

Economic Variables

GDP growth
Unemployment rate*
Employment growth

* Annual change in unemployment rate
Sources: OECD; RBA

Figure A4: France data

Financial Variables

* Annual change in bond rates
Sources: BIS; OECD; RBA

Economic Variables

* Annual change in unemployment rate
Sources: OECD; RBA

Figure A5: Germany data

Financial Variables

* Annual change in bond rates
Sources: BIS; OECD; RBA

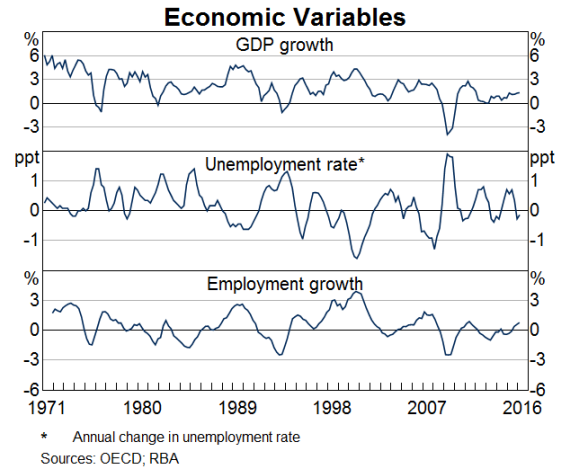## Economic Variables

* Annual change in unemployment rate
Sources: OECD; RBA
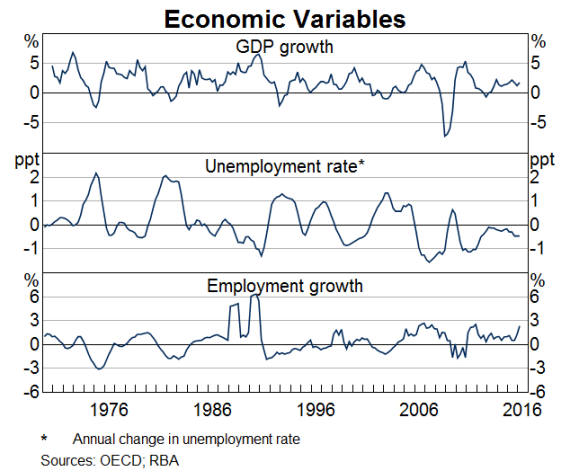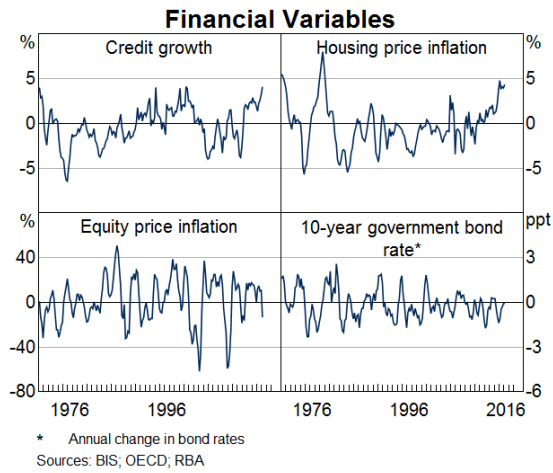
Appendix B: Cycles in stocks and flows

We assume a continuous time model. All debt is taken out at time τ is fully paid-off by period τ+T (i.e. it is T-period debt). Let $D_\tau(t)$ be the debt outstanding at time t for debt drawn down at time τ. For credit-foncier debt, the general expression for the outstanding balance of debt at time t is given by

$$D_\tau(t) = D_\tau(\tau)\left(e^{R(t-\tau)} - \frac{e^{RT}\left(e^{R(t-\tau)} - 1\right)}{e^{RT} - 1}\right) \quad \text{if } \tau \leq t \leq \tau + T$$

$$= 0 \text{ otherwise}$$

At time t, the total outstanding stock of debt S(t) is equal to the sum of all outstanding debt drawn down between period t-T and t

$$S(t) = \int_{t-T}^{t} D_\tau(t)d\tau$$

Now, suppose the flow of new debt, $D_\tau(\tau)$ in period τ has a cycle. So let $D_\tau(\tau) = A\sin(\alpha\tau) + B$ (note that there is no phase shift, without loss of generality).

Then it can be shown that

$$S(t) = -A\frac{e^{Rt}}{e^{RT} - 1}\left(\frac{e^{-Rt}}{R^2 - \alpha^2}\right)\left(-R\sin(\alpha t) - \alpha\cos(\alpha t) + Re^{Rt}\left(\sin(\alpha(t-T)) - \alpha e^{RT}\cos(\alpha(t-T))\right)\right)$$

$$+ \frac{A}{\alpha}\frac{e^{RT}}{e^{RT} - 1}(\cos(\alpha(t-T)) - \cos(\alpha t)) - \frac{B}{R} + BT\frac{e^{RT}}{e^{RT} - 1}$$

So S(t) has frequency $\frac{2\pi}{\alpha}$, albeit with a phase shift.

Details on the derivations are available from the authors upon request.

Appendix C: Robustness checks

Table C2: Estimates of financial and business cycles – robustness check[a]

Peaks in multivariate spectra

	Financial cycle estimate	Confidence interval	Business cycle estimate	Confidence interval
	Years	95%	Years	95%
US	8.1	[6.0, 36.3]	8.1	[4.3, 14.5]
Australia	8.1	[3.2, 24.2]	4.3	[3.5, 14.5]
UK	12.1	[4.5, 24.2]	8.1	[4.8, 18.1]
France	10.4	[5.2, 24.2]	9.1	[4.5, 72.5]
Germany	8.1	[3.2,14.5]	10.4	[3.8, 18.1]

(a) Sample size limited to 1980:1 to 2016:1 for all countries
Sources: Bank of England; BIS; INSEE; OECD; RBA

Table C3: Estimates of financial and business cycles – robustness check[a]

Peaks in multivariate spectra

	Financial cycle estimate	Confidence interval	Business cycle estimate	Confidence interval
	Years	95%	Years	95%
US	12.9	[7.5, 22.6]	5.7	[4.5, 11.3]
Australia	14.5	[4.3,18.1]	4.3	[3.6, 10.4]
UK	13.4	[4.7, 23.5]	8.5	[4.5, 15.7]
France	12.6	[4.9, 17.7]	9.8	[3.1, 14.8]
Germany	12.9	[7.5,22.6]	9.1	[4.7, 18.1]

(a) Only using credit growth and housing price inflation as financial variables
Sources: Bank of England; BIS; INSEE; OECD; RBA

Table C4: Estimates of financial and business cycles– robustness check[a]

Average cycle lengths using Bry-Boschan quarterly algorithm

	Financial variables[b]			Economic variables[c]		
	Contraction	Expansion	Cycle	Contraction	Expansion	Cycle
US	2.3	9.6	11.5	1.2	4.7	6.4
Australia	3.1	2.5	5.6	1.8	7.4	9.8
UK	6.4	5.6	8.2	1.8	5.2	7.0
France	3.8	4.9	9.8	1.6	6.0	8.5
Germany	2.8	0.5	3.3	2.0	2.9	5.1

(a) Only using credit growth and housing price inflation as financial variables
(b) Includes credit, housing prices
(c) Includes GDP, employment and unemployment rate
Sources: Bank of England; BIS; INSEE; OECD; RBA

Table C5: Estimates of financial and business cycles – robustness check[a]

Peaks in multivariate spectra

	Financial cycle estimate	Confidence interval	Business cycle estimate[b]	Confidence interval
	Years	95%	Years	95%
US	7.9	[5.2, 49.5]	6.2	[5.0, 19.8]
Australia	4.2	[2.6, 42.0]	6.7	[2.2, 30.3]
UK	13.0	[4.2, 18.2]	7.6	[2.5, 91.0]
France	14.1	[6.3, 56.5]	12.6	[2.6, 56.5]
Germany	12.7	[2.0, 29.7]	12.7	[3.0, 29.7]

(a) Using the annual Jordà, Schularick and Taylor (forthcoming) database, using data from 1920 onwards
(b) The only economic variable available in this database is GDP growth
Sources: Jordà, Schularick and Taylor (forthcoming); RBA

# Appendix D: Syncrhonisation of cycles across countries

Table D1 suggests that the degree of synchronisation in longer cycles in credit growth has been highest between the US, the UK and Australia. For longer cycles in GDP growth, the degree of synchronisation has been highest between the US and the UK, and between France and Germany. These results are largely to be expected. The connections between the UK and the US are likely to reflect strong financial and economic linkages. The results for the France and Germany are likely to reflect the presence of strong regional economic and financial linkages as well as common exposures to other European countries. Being a small, open, resource-based economy, Australia's longer cycle in GDP growth does not seem to be substantially linked to those in the larger advanced economies; on the other hand, the longer cycle in credit growth appears to have been highly connected with the credit cycles in the US and the UK. These results highlight that in addition to domestic conditions, conditions in other countries might also matter for the financial and business cycles.

Table D1: Synchronisation in longer cycles across countries

Concordance statistics using Bry-Boschan quarterly algorithm[a]

Credit growth	GDP growth				
	US	Australia	UK	France	Germany
US		0.65	0.82	0.36	0.50
Australia	0.85		0.62	0.45	0.57
UK	0.73	0.87		0.52	0.68
France	0.34	0.50	0.61		0.83
Germany	0.39	0.43	0.50	0.60	

(a) Concordance statistics represent the proportion of time the cycles are in the same phase
Sources: BIS; OECD; RBA

Figure D1
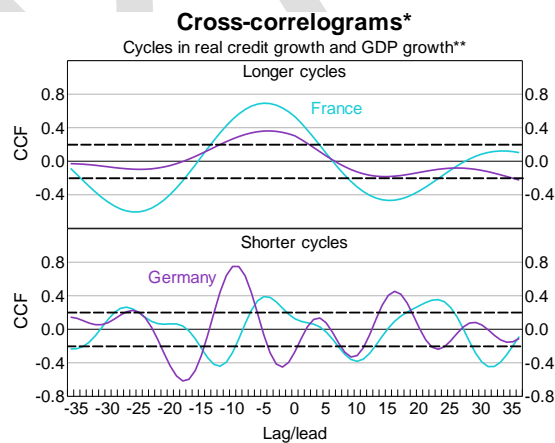
Cross-correlograms*

Cycles in real credit growth and GDP growth**

*   Shows the correlation between the cycle in credit growth at time 0 with 36 leads and lags of the cycle in GDP growth; dashed lines represent approximate 95% confidence intervals
**   Longer cycles are between 8 and 30 years; shorter cycles are between 2.5 and 8 years
Sources: BIS; OECD; RBA

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