Dollar Reserves and U.S. Yields: Identifying the Price Impact of Official Flows

by Rashad Ahmed and Alessandro Rebucci

Discussion at the 2023 ABFER Conference

Martin Berka

Massey University, ABFER, CAMA

May 23, 2023

▲□▶ ▲□▶ ▲□▶ ▲□▶ ▲□ ● ● ●

Interesting and timely contribution

- Impact of foreign official (FO) flows on US 5Y, 10Y, and 30Y yields is more than twice previously estimated (1999M1 - 2018M12)
 - On impact: FO sale of \$100bn raises yields by > 100bp, (19-44bp in lit)
 - If identified in SVAR through heteroskedasticity
 - Include 'omitted variables', i.a., foreign g't yields, and FRB shocks (Swanson 2021)
 - Robust to including Private flows and International yield factors
- IRFs are significant at 10% (for a while)
- Doomsday implication:
 - Shift by China's UST holdings by 1% raises yields by 24.4bp !

- Authors note global economic conditions a key driver ("precautionary, mercantilist, and exchange rate smoothing motives" by other countries)
 - Should you include OECD (non-US) GDP series as controls?
- Results with foreign yields (& structural budget positions) should be the benchmark rather than robustness?
- More generally: bilateral panel VAR?
 - China vs Saudi Arabia exercise shows how heterogeneous circumstances drive FO flows

• Get an FO flows panel even for a subset of countries to assess the importance of partner conditions

• I am confused by the early discussion of 'simultaneity bias', which you say arises from:

▲□▶▲□▶▲≡▶▲≡▶ ≡ めぬぐ

• Causal relation (a) \Uparrow Demand for UST $\Rightarrow \Downarrow$ yield $[\mathit{cov}(\mathit{UST}, \mathit{yield}) < 0]$

- I am confused by the early discussion of 'simultaneity bias', which you say arises from:
 - Causal relation (a) \uparrow Demand for UST $\Rightarrow \Downarrow$ yield [cov(UST, yield) < 0]
 - Causal relation (b) \Downarrow yield \Rightarrow \Downarrow Demand [cov(UST, yield) > 0]

▲□▶▲□▶▲≡▶▲≡▶ ≡ めぬぐ

- I am confused by the early discussion of 'simultaneity bias', which you say arises from:
 - Causal relation (a) \uparrow Demand for UST $\Rightarrow \Downarrow$ yield [cov(UST, yield) < 0]
 - Causal relation (b) \Downarrow yield $\Rightarrow \Downarrow$ Demand [*cov*(*UST*, *yield*) > 0]
 - "... an estimate which confounds (a) and (b) will be less negative than the true causal effect (a) which we wish to estimate"

- Isn't this just description of the effects of shift in demand? In a D& S framework?
 - discuss supply-factor for identification



▲□▶ ▲□▶ ▲ 三▶ ▲ 三▶ 三三 - のへぐ

Many international drivers of UST demand

• "precautionary, mercantilist, and exchange rate smoothing motives"

- but only control for global yields (common factor)
- Plenty to add?
 - global GDP growth (monthly economic activity indices)
 - global financial conditions
 - global structural deficits
 - EER
 - panel VAR

• Brunnermeier et al. (2021), Lutkepohl et al. (2021), etc.

$$\begin{array}{rcl} \mathsf{E}(u_t u_t') &=& \Sigma_i, \ i \in \{1,2\}, \Sigma_1 \neq \Sigma_2 \\ \Sigma_1 &=& BB' \\ \Sigma_2 &=& B \Lambda B' \end{array}$$

- Λ diagonal, λ_{kk} > 0, and distinct (error variance isn't scaled equally across variables)
- It is not the change of volatility in macro aggregates that is in doubt (GFC!), instead:
 - the shock variances change in a way that is distinct across variables (needed for uniqueness of B)
 - 2 the responses of variables (pre/post Lehman) are unchanged
- Then, structural shocks can be recovered

• The exposition could be clarified. Authors show:

- Discussion why Lehman Brothers the right place for variance change
- Point to other studies
- Test for a single known structural break in Sep 2008
- 3 "unknown-date" test breaks: April 03, May 08, May 11
- Testing with FO^2 produces 3 different (though similar) breakpoints
- Provide evidence of VAR parameter stability (in estimated SVAR)
- Tests to support the assumption that λ_{kk} are distinct
- Too many steps? I would like to see that the break tests are there to illustrate, not to guide the assumed breakpoint test, maybe move the rest to an appendix?



Figure 3: Absolute FO Flows and U.S. Dollar Returns

The figure plots absolute values of FO flows and monthly log-returns of the trade-weighted U.S. Dollar in the left and right panels, respectively. The thick dashed vertical line is September 2008. The thin dot-dashed vertical lines are detected breaks in April 2003, May 2008, and May 2011 in the absolute FO flows series using the testing framework in Bai and Perron [2003]. For absolute FO flows (absolute Dollar returns), the pre September 2008 mean is 0.15 (0.009) and the post September 2008 mean is 0.18 (0.127).

- 3 tested breaks visible
- "High volatility" period seems short (May 08 May 11)
- Followed by period with less FO volatility than pre-GFC?

- Bai & Perron (2003) multiple breaks test is for univariate series: discuss how you applied in VAR
- Clearly state which breakpoint date is used in the SVAR
- You test $\lambda_i \neq \lambda_j$, but your identification also requires $\lambda_i > 1 \ \forall i$. Report the estimated λ_i , $i = \{1:6\}$
 - The other conditions are 'necessary but not sufficient' for your story to hold

- Although you identify off a break in FO flows, over the sample these go from 20% to 30% of total flows (with a spike around GFC)
 - FP flows can have different drivers but this is good for your identification of FO flows' impact
 - FP flows can have the same drivers large financial institutions can dwarf some CBs
- FP flows should be in your benchmark both to avoid omitted variable critique, and to help identification
- It'd be illustrative to see responses to FP (as opposed to FO) shocks

- Use panel, or at least help identification using a few bilateral VARs to illustrate
 - Just like with FP flows, Panel gives you the country-level heterogeneity of *reasons* for FO flows
 - At least focus on large players, notably China, G7, etc.
 - Not about 'who' is selling but 'why' they're selling reasons differ and this can affect your elasticity

・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・
・

• Your application of global SVAR estimates to China or Saudi Arabia is internally inconsistent: you note heterogeneous reasons yet assume homogeneity

- "Twice as big" a good sales pitch but invites scrutiny
- You argue this is because simultaneity bias / endogeneity, and OVB
- I think your estimation points to evolving FO flow drivers pre/post GFC, and that these translate to different supply elasticities (post-GFC, S relatively less elastic)
 - Swanson (2021) shocks help identify supply side changes estimate your SVAR without them and see if magnitude is still double
- Estimate in split samples (pre- vs post- GFC, or within the break points you identify) to see how IRFs change?
- Presumably significant differences in IRFs, and hence: why? What is it about GFC that caused the FO volatility change

- Reasons for starting the sample in 1999?
- I'd welcome clearer links to international literature, e.g.:

▲□▶ ▲□▶ ▲ 三▶ ▲ 三▶ 三 のへぐ

- GFC(ycle)
- USD dominance
- Safe assets
- International capital flows more general

Thank you